Contents lists available at ScienceDirect

Journal of Financial Economics

journal homepage: www.elsevier.com/locate/jfec



Benjamin Balsmeier^a, Lee Fleming^{b,c}, Gustavo Manso^{c,*}

^a KOF Swiss Economic Institute, ETH Zurich, Leonhardstrasse 21, Zurich 8092, Switzerland

^b Fung Institute for Engineering Leadership and Haas School of Business, University of California, 2451 Ridge Road, Berkeley 94709, USA

^c Haas School of Business, University of California, 545 Student Services Building 1900, Berkeley 94720, USA

ARTICLE INFO

Article history: Received 16 October 2015 Revised 12 April 2016 Accepted 18 April 2016 Available online 21 December 2016

JEL Classification: G34

L14 L25

M21

Keywords: Corporate governance Board composition Innovation Exploration and exploitation

1. Introduction

The board of directors has an important role in the governance of corporations. Charged with overseeing and advising managers, it can effectively reduce agency costs that arise from the separation of ownership and control. Several authors have argued that independent directors, with no ties to the company other than their directorship, are better suited to perform this role as they can credibly limit managerial discretion by punishing managers

* Corresponding author.

E-mail address: manso@berkeley.edu (G. Manso).

http://dx.doi.org/10.1016/j.jfineco.2016.12.005 0304-405X/© 2016 Elsevier B.V. All rights reserved.

ABSTRACT

Much research has suggested that independent boards of directors are more effective in reducing agency costs and improving firm governance. How they influence innovation is less clear. Relying on regulatory changes, we show that firms that transition to independent boards focus on more crowded and familiar areas of technology. They patent and claim more and receive more total future citations to their patents. However, the citation increase comes mainly from incremental patents in the middle of the citation distribution; the numbers of uncited and highly cited patents—arguably associated with riskier innovation strategies—do not change significantly.

© 2016 Elsevier B.V. All rights reserved.

after undesirable outcomes (e.g. Fama and Jensen, 1983; Williamson, 1983).

We investigate the effect of board independence on search and innovation processes. Relying on regulatory changes for identification, we show that firms that transition to independent boards patent and claim more and that their patents receive more citations. However, these patents are in crowded and familiar areas of technology and fall into the middle of the citation distribution.

The patent and citation count increase is consistent with many classical models. Independent boards are more likely to terminate the manager in case of poor performance (Weisbach, 1988) and this threat provides an incentive to the manager to work hard (Stiglitz and Weiss, 1983). Increased monitoring from independent boards may alleviate agency problems such as shirking or tunneling of corporate resources. Managers should also take actions that are—and appear to be—closer to the interests of shareholders (Harris and Raviv, 1978; Holmstrom, 1979; Holmstrom and Milgrom, 1991). When under increased scrutiny





CrossMark

^{*} The authors thank David Hsu, Adair Morse, David Sraer for helpful comments, as well as Gabe Fierro and Guan-Cheng Li for invaluable research assistance. Balsmeier acknowledges financial support from the Flemish Science Foundation. This work is supported by NSF grant 1360228 and the Coleman Fung Institute for Engineering Leadership; errors and omissions remain the authors'.

and demands for results, managers will focus on quantifiable results, such as a greater number of patents. They will adduce an increase in patent count to satisfy demands for performance.

Innovation strategies are more complex, however, than what is reflected in simple patent and citation counts. We distinguish between exploration of new technologies and exploitation of well-known technologies (March, 1991). On one hand, independent boards may provide greater diversity of opinion and expertise outside the manager's competence that enhance exploration. On the other hand, independent boards may cause a manager to focus on exploitation in order to maximize the mean outcome, instead of exploration which could open up breakthroughs. Because independent boards are more likely to fire a manager after poor performance, managers will tend to pursue less exploratory projects (Manso, 2011). Managers may avoid new technologies that might be construed as empire building (Jensen, 1986). Boards may also directly resist exploration of new areas, if they fear that in the short-term the stock market fails to properly value investments in innovation (Stein, 1989; Cohen, Diether, and Malloy, 2013). Due to potential conflict of interests between independent board members and the manager, or alternately, less familiarity with the firm's industry and technology, the quality of research advice given by independent boards may be different (Adams and Ferreira, 2007).

These arguments imply observable outcomes. Firms whose boards become independent will patent and claim more but the increase will come mainly from patents in areas the firm has previously patented in; the effect on patenting in new areas is less clear. Citations to the firm's patents will increase, but the citation increase will come mainly from patents in the middle of the citation distribution; the effect on risky technologies that might provide a breakthrough or fail completely is similarly unclear. Furthermore, the increase in citations in the middle of the distribution will be mediated by the movement of the firm into better known and more crowded areas of technical search; in other words, the firms' patents will be more highly cited simply as an artifact of their search strategy and the citation norms of more crowded fields. Managers will exert effort towards maximizing the return of innovation on previously proven trajectories, but may invest less in new technological trajectories.

Evidence comes from observing search strategies for firms that were forced by regulatory changes to adopt more independent boards. Starting in 2002, stock exchanges and the Sarbanes-Oxley Act (SOX) required firms to have a majority of independent directors (for a similar approach, see Duchin, Matsusaka, and Oguzhan (2010)). Comparing firms that changed from less to more independent boards against firms that already had independent boards, we find increased innovation activity overall, but no significant effect on more explorative types of innovation. Firms whose boards become more independent patent more and receive more citations to their patents, however, the effects are not significant for uncited and highly cited patents. Firms whose boards become more independent also work in more crowded and more familiar technologies; the rates of prior and self-citation increase. Moreover, results are more pronounced for firms with high Research & Development R&D stock and a high entrenchment index.

The effect of board independence on innovation is immediate, taking place only one year after the transition to an independent board. This is surprisingly quick, as research takes time: programs must be funded, staffed, and executed, after which any successful results must be patented. It is possible though less likely that a new R&D strategy could effect such immediate impact; a more plausible explanation is that the patenting processes changed, and in particular, that the firm's engineers and lawyers looked more carefully for patentable technology within the firm's extant portfolio. This would be reflected in the immediate increase of patents, claims, and even citations.

The estimated economic impact of board independence appears large. Depending on the model specification, the number of patents increases between 20% and 30%, and the number of citations between 40% and 60%. In absolute terms, however, the median number of forward cites is five, which implies three more future citations. Similarly, as firms move into more crowded technological areas, as indicated by more backward citations, their chances of receiving citations increases independent of the guality of the innovation, simply because more competing firms/inventors work on similar areas and thus cite each other. Backward citations increase by up to 49.8%, which translates into approximately 13 more backward citations in absolute terms for the median firm. Altogether, the evidence supports arguments for a more nuanced relationship between oversight and innovation; greater oversight appears to lead to greater focus and productivity but have no impact on breakthroughs.

2. Literature review

A large literature studies the role and influence of board characteristics (for an overview see Adams, Hermalin, and Weisbach (2010); for the economic relevance of boards see Ahern and Dittmar (2012)). Much of the literature focuses on the role of independent board members (most recently, e.g., Masulis and Mobbs, 2014; Brochet and Srinivasan, 2014). Several studies have analyzed how independent directors influence Chief Executive Officer (CEO) compensation (e.g., Faleye, Hoitash, and Hoitash, 2011; Coles, Daniel, and Naveen, 2008; Denis and Sarin, 1999; Core, Holthausen, and Larcker, 1999), CEO appointments and dismissals (Knyazeva, Knyazeva, and Masulis, 2013; Borokhovich, Parrino, and Trapani, 1996; Weisbach, 1988), adoption of antitakeover defenses (Brickley, Coles, and Terry, 1994) or takeover premiums (Cotter, Shivdasani, and Zenner, 1997; Byrd and Hickman, 1992). From these studies the picture emerges that independent board members increase board oversight. Whether such intensified board monitoring is beneficial or detrimental to shareholder wealth is less clear and may depend on the complexity of a firm's operations (Faleye, Hoitash, and Hoitash, 2011; Duchin, Matsusaka, and Oguzhan, 2010).

Several recent papers use patent data to empirically study how corporate governance affects innovation. Raw patent counts are usually supplemented by the number of

citations that a patent receives, as this measure correlates with financial and technical value (Harhoff, Narin, Scherer, and Vopel, 1999; Hall, Jaffe, and Trajtenberg, 2005); future cites are sometimes broken down by whether the firm cites its own work (Lerner, Sorensen, and Stromberg, 2011). Though less common, some papers have analyzed technology classes or the tails of the citation distribution (Gonzalez-Uribe and Xu, 2016; Byun, Oh, and Xia, 2015; Cerqueiro, Hegde, Penas, and Seamans, 2015; Chen, Gao, Hsu, and Li, 2015). Measures of originality and generality (Hall, Jaffe, and Traitenberg, 2001) have also been used, though these measures depend on the US Patent and Trademark Office's changing and now discontinued classification of technologies (see Lerner, Sorensen, and Stromberg, 2011; and Hsu, Tian, and Xu, 2014). Lerner and Seru (2014) detail a number of problems with the use of patent measures in the finance literature, including failures to correct for differences in time periods and truncation (caused by the lag between application and patent grant, or delay in the accumulation of future prior art citations), economic value (typically proxied by future prior art citation), technology (typically measured by the United States Patent and Trademark Office classes), and disambiguation of assignees (determining which firms own which patents). One empirical contribution of this paper is to offer improved and easily calculated measures that can address some of these issues.

The results of the recent surge of empirical patent work on governance and innovation are decidedly mixed. Much of the contradictory work is well identified, so resolution will have to rely on sharper theory or more careful measurements of governance and innovation. A variety of papers find that stronger governance leads to greater innovation (alternately, weaker governance leads to decreased innovation). Aghion, van Reenen, and Zingales (2013) show that greater institutional ownership correlates with greater patenting and citations to patents. Bernstein (2014) finds that firms experience no change in the amount of patenting following an Initial Public Offering (IPO) (when they would assumedly transition from strong oversight by venture capitalists to weaker public oversight), however, they do experience a decrease in citations. Atanassov (2013) found that a strengthening of antitakeover provisions in a state (assumedly implying weaker governance) led to fewer patents and citations, but that institutional shareholders decreased the effect. Sapra. Subramanian, and Subramanian (2014) used a similar research context to Atanassov (2013) but found a non-monotonic effect, where innovation increased for firms that experienced very weak and very strong external takeover pressure.

In contrast, a variety of papers finds that weaker governance leads to increased innovation (alternately, stronger governance leads to decreased innovation). Atanassov (2016) finds that firms with a greater proportion of bank financing invented more and more highly cited patents (and that the volatility of citations was greater). In contrast to Atanassov (2013) and in partial contrast to Sapra, Subramanian, and Subramanian (2014), Chemmanur and Tian (2016) find that firms with greater anti-takeover provisions receive more and more highly cited patents. Most similar to the current study, Faleye, Hoitash, and Hoitash (2011) find that monitoring intensity, as measured by the proportion of independent board directors on at least two monitoring committees, correlates negatively with research and development expenditure and future prior art citation counts. While they present wellspecified panel data models, their Sarbanes-Oxley regressions (the main identification strategy used in the current paper) investigate the effect of SOX on firm value but not, however, on R&D and patent data. Kang, Liu, Low, and Zhang (2014) find no correlation with social connections between the CEO and board members and research and development spending (arguably a social connection implies weaker governance); they find a positive correlation with patents and citations.

Using a differences-in-differences (DiD) identification strategy based on the regulatory requirements of SOX, we find no effect of a firm's transition to an independent board upon R&D spending but a positive effect on total patenting and citations and a focusing of the firm's innovative search on known and previously successful areas; these results remain robust across a variety of matched, fixed effects, and trend control models.

Taking heed of the critiques of Lerner and Seru (2014), this paper assembles a suite of more detailed and nuanced measures of innovation. This battery of measures offers additional and consistent insights into the mechanisms of how board independence influences innovation, while retaining the advantages of the SOX identification strategy. Of more general interest, the battery of measures enables cleaner identification of a firm's search strategy; it illustrates how a firm can invent more highly cited and valuable patents by exploiting its current area of expertise. Such exploitation may be characterized as local search and is probably less risky, yet it is still innovation and arguably the most effective and valuable search strategy for the firm.

3. Identification strategy

Identification for our study relies upon regulatory changes that forced public firms to increase the presence of independent directors on their boards in the early 2000s. The effects of those regulatory changes on variables other than innovation have been analyzed elsewhere (see, e.g., Duchin, Matsusaka, and Oguzhan (2010), for a setup that is most similar to ours). In this section, we briefly describe the regulatory framework that is relevant to our analysis.

Initiated by recommendations of the Blue Ribbon Committee (BRC) in 1999, stock market rules of the NYSE and Nasdaq have been built upon the assumption that independent board members are better able to monitor managers. Subsequent to the BRC recommendations, the Securities and Exchange Commission (SEC) approved new rules in December 1999, requiring public firms to move to a fully independent audit committee with the next re-election or replacement of audit committee members. Further motivated by prominent corporate scandals, e.g., Enron, this rule was written into U.S. law in 2002 as a part of the Sarbanes-Oxley Act (SOX). It was followed by subsequent



Fig. 1. Fractions of independent boards and directors. This graph illustrates the evolution of independent boards over the sampling period. A board is defined as independent in the empirical estimations if the majority of board members are classified as independent by the Investor Responsibility Research Center (IRRC). In this graph, independent directors represents the average fraction of independent board members of all firms in the study. Descriptive statistics are shown in Table 1.

NYSE and Nasdaq regulations in 2003 that imposed even stricter requirements on board composition. In addition to having an audit committee composed of exclusively independent directors, both stock exchanges forced firms to have a majority of independent directors as regular board members, and the compensation and nomination committees had to consist of 100% independent board members (> 50% if firms are listed on Nasdaq only).

Definitions of director independence vary slightly across each rule. SOX states in Section 301 that a given director is independent if the person does not "accept any consulting, advisory, or other compensatory fee from the issuer" (except for serving the board), and is not an "affiliated person of the issuer or any subsidiary" (NYSE speaks of "no material relationship"; and Nasdaq requires no relationship that would interfere with "independent judgment"). The NYSE and Nasdaq regulations are clear; the independence assumption is violated, for instance, if a director him- or herself or a direct family member was an employee of the firm during the previous three years, or a family member works for a third firm with which the given firm has a professional relationship, or a family member is connected to the firm's auditor.

These regulations made board changes necessary for a large group of firms. The number and fraction of independent board members was fairly stable until the year 2000. As the described board regulations came into effect, more and more independent directors were appointed to corporate boards. Fig. 1 illustrates the changes in board composition for the sample of firms used in our study. It resembles a pattern that has been documented in other studies for differing sets of public firms (e.g., Linck, Netter, and Yang, 2008; Duchin, Matsusaka, and Oguzhan, 2010). Board composition data are taken from the Investor Responsibility Research Center (IRRC). From 1996 to 2006, the IRRC tracked individual board members of all major public U.S. firms and indicated in their database whether an individual board member is independent, an employee of the firm, or otherwise affiliated (former employee, employee of an organization that receives charitable gifts from the company, employee of a customer or supplier to the company, relative of an executive director, etc.).

Reflecting the previously introduced regulatory changes, Fig. 1 shows an increase of independent director presence on corporate boards from 2001 to 2006. Theoretical considerations about board control suggest that a crucial difference arises when a board switches from a minority to a majority of independent board members (Harris and Raviv, 2008).¹ It was further an explicit requirement of regulatory reforms. Thus, our analysis focuses on this variable. Our data also show that the proportion of firms with a majority of independent board members stayed rather stable around 68% before 2000 and moved up to about 94% by 2006.

Our empirical identification of the relationship between board independence and innovation stems from the difference between firms who were already in compliance with the regulatory changes before 2001 and those firms who switch to a majority of independent directors after regulatory changes became effective. Hence, all firms that were not required to change their board serve as a control group. In line with Duchin, Matsusaka, and Oguzhan (2010), we define firms as treated when they switch to an independent board in 2001 or later and have an audit committee that contains 100% independent board members. The latter requirement helps to sort out potential voluntary switches, increasing the amount of truly exogenous increases of independent board members and making our main variable of interest less likely to be confounded by endogenous choice. The fraction of independent directors increased by 25% during 2001-2006 within noncompliant firms and by 9% within firms that had already fulfilled the regulatory requirements before 2001.

4. Sample selection

The data set we built up for our study is determined by the joint availability of data on the composition of corporate boards and committees from the IRRC, information on basic firm characteristics from Compustat, and patent data from the National Bureau of Economic Research (NBER), the Fung Institute, and the United States Patent and Trademark Office (USPTO). The IRRC provides data on corporate board members for 3,000 major public U.S.-based firms from 1996 to 2006. Compustat has further information on almost all of the firms covered by IRRC. A major challenge for the empirical researcher interested in those firms' innovative activities is the identification and compilation of the corresponding patent portfolios. Researchers involved in the NBER patent data project have spent significant

¹ The fraction of independent board members provides more variation but has two major disadvantages. First, considering board voting behavior, it is likely that the influence of independent directors on board oversight does not increase linearly with the number or fraction of independent members but exhibits a jump when independent directors gain or lose the majority of votes. Second, the switch from a minority to a majority of independent directors was an explicit requirement of regulation, such that it is more likely that observed changes in that regard happened involuntarily, which in turn improves the identification of causal effects.

540

amounts of resources to identify patents that have been granted to U.S.-based firms. The NBER patent database contains, however, only those patents that have been granted through 2006. Due to the time lag with which inventions are granted property rights (1-5 years) and the publication of corresponding data by the USPTO, this results in significantly truncated data for patents filed after 2001. Researchers have found ways to use incomplete patent data for the years 2002–2006, exploiting the distribution of applications before 2002, but those approaches add noise to econometric analyses, and lead to significant estimation errors in our case, because our sample of board data covers 50% of years for which the NBER data are severely truncated. The issue becomes even more prevalent if researchers want to take citations to patents into account that often occur several years after a patent has been granted. In terms of patent applications, the NBER data misses 18% of patent applications of U.S.-based assignees identified in 2002, rising to 99% by 2006.²

Newly available disambiguations (see Balsmeier, Chesebro, Fierro, Johnson, Kaulagi, Li, Lueck, O'Reagan, Yeh, Zang, and Fleming, 2016) provide more recent data, avoid the truncation of the NBER patent database, and identify comprehensive patent portfolios of the firms in our sample up to the year 2007.³ Following the literature (e.g., He and Tian, 2013), we assign an eventually granted patent to the year it was applied for. Disambiguation of firm names presents a major challenge, since patent documents do not contain a unique identifier of assignees. Following disambiguation, patents are aggregated to the firm level and merged with other databases such as Compustat and IRRC.

We extended the reach of the NBER patent database by combining it with USPTO and Fung Institute data, including patent citations and other detailed information within each patent document. We started with standardized assignee names provided by the USPTO for all patents granted through December 31, 2012. These standardized assignee names are largely free of misspellings but still contain many name abbreviations for individual firms. The standardized USPTO assignee names remain consistent throughout time and have been used by the NBER patent project team to disambiguate firm names. For almost all U.S. firms that received at least one patent between 1975 and 2006, the NBER provides a unique time-invariant assignee. We took all variations of standardized assignee names that belong to a given single firm as a training set. and gave all granted patents that appear with the same standardized assignee name the same unique NBER identifier.⁴ This information enabled us to track firms' patenting activity over significantly longer time periods, overcoming truncation issues of patent applications and generally increasing the accuracy of available patent portfolios.

Finally, we merged unique time-invariant Compustat identifiers to the patent assignee identifiers as they are provided by the NBER. It is worthwhile to note that in our analysis we take only those firms into account for which the NBER has identified Compustat matches, and we assigned zero patents only to those firms where the NBER team searched for but could not find matches with any patent. In this regard we deviate from other studies that assign zero patents also to those firms that have not been tested to appear as a patent assignee or not. Thus, we avoid measurement errors at the expense of a smaller but more accurate data set.

In order to circumvent potential selection effects to confound our estimation of the relationship between board independence and innovation, we further removed all firms that appear only before the year 2000 or entered the sample in the year 2000 or later, such that the remaining firms can be observed over a timespan where the previously described regulatory changes took place. Finally, we arrive at a sample of 6,107 observations on 713 firms observed during the period from 1996 to 2006 for which we could gather all information of interest. All firms in the sample combined have applied for and been granted 328,463 patents during the sample period.

4.1. Measuring innovative search

Much recent empirical work on corporate governance and innovation has relied on patent data (e.g., Atanassov, 2013; He and Tian, 2013). Raw patent counts are used as well as the number of future prior art citations that a patent receives, as the number of future cites correlates with financial and technical value; highly cited patents are much more valuable commercially and the relationship is highly skewed in favor of very highly cited patents (Harhoff, Narin, Scherer, and Vopel, 1999); (Hall, Jaffe, and Trajtenberg, 2005). To be comparable with the extant literature we will show how board independence influences patent counts and citations. Our results go on to illustrate, however, that raw patent counts and total citation counts are of limited use in identifying differences in innovative search strategies, specifically towards more or less exploration. Therefore, we introduce a suite of measures, consistent with the arguments of Lanjouw and Schankerman (2004) for the use of multiple indicators of patent quality. These serve as additional dependent variables besides raw patent counts and citations, thus enabling the illustration of a richer and more robust picture of how board independence affects not only the rate but also the type and direction of innovation.

First, we calculate the number of citations that each patent makes to prior patents (Lanjouw and Schankerman, 2004). An increase in the number of backward citations reflects direct relations to more prior art that must be specified in the patent application (required by law). This correlates with innovative search in relatively more crowded, better-known, and typically more mature technological areas.

² The numbers are derived by comparing all patent applications in the NBER database with all patents in the Fung Institutes database as published in April 2014.

³ We gather patent data through 2007, because we will estimate regressions of firms' patenting activities in year t on board data and controls in t-1, reflecting that patenting activities need some time to be influenced by boards and simultaneous determination of variables may otherwise confound the estimation.

⁴ Based on the first assignee that appears on the patent document. It allowed us to identify "250k additional patents granted to U.S.-based assignees after 2006.

Second, we take the number of times a given patent cites other patents owned by the same company (Sorensen and Stuart (2000); similar measures are used in Faleye, Hoitash, and Hoitash (2011)). More self-cites indicate search within previously known areas of expertise while fewer self-citations indicate a broadening of innovative search or efforts to explore areas that are new to the firm.

Third, we calculate the number of patents that are filed in technology classes previously unknown to the firm. Unknown patent classes are defined as those in which a given firm has not applied for any patent beforehand (starting in 1976). The complement is the number of patents applied for in known classes. Addressing one concern of Lerner and Seru (2014), we consistently use the original patent class at time of patent grant; hence, if the USPTO defines a brand new class and issues a new patent, it will be observed, but if the USPTO redefines an old patent into a new class, it will not change the measure.

A continuous measure of whether firms stay or deviate from known research areas is the technological proximity between the patents filed in year t and the existing patent portfolio held by the same firm up to year t-1 (Jaffe, 1989):

$$P_{it} = \sum_{k=1}^{K} f_{ikt} f_{ikt_{-1}} / \left(\sum_{k=1}^{K} f_{ikt}^2 \cdot \sum_{k=1}^{K} f_{ikt_{-1}}^2 \right)^{\frac{1}{2}}$$

where f_{ikt} is the fraction of firm *i*'s patents that belong to patent class *k* at time*t*, and $f_{ikt_{-1}}$ is the fraction of firm *i*'s patent portfolio up to *t*-1 that belongs to patent class *k*. P_{it} ranges between zero and one. The highest possible value indicates that the patents filed in year *t* are distributed across patent classes in the exact same way as the portfolio of all patents of the same firm up to the previous year.⁵ Positive coefficients in a regression would thus indicate a more narrow innovation trajectory within known areas.

Fourth, we categorize patents according to how many citations they have received relative to other granted patents that have been applied for in the same technology class and year (Azoulay, Graff Zivin, and Manso, 2011). In addition to limiting comparison of similar patents, we exclusively and exhaustively bin all patents according to their location in the distribution of citations. This is intended to clearly separate different types and degrees of innovative outcomes, ranging from highly successful breakthroughs (highly cited) to completely failed inventions (not cited at all) and moderately successful outcomes that lie between. We estimate separate models for each of the four non-overlapping categories: top 1%, 2nd-10th%, not in the top 10% but cited at least once, and never cited at all. We count a patent as a top 1% (2-10%) patent if the patent falls into the highest percentile (centile) of the citation distribution in the same technology class and application year. We also separately count all patents that received no citation at all and those that have received at least one citation but do not fall in the top 10% category.

Fifth and finally, we calculate the total number of claims made by a firm's patent portfolio each year (Lanjouw and Schankerman, 2004). It is difficult to algorithmically interpret ex ante the innovative value of any particular claim, however, as claims can be added as scope conditions which typically act as limitations on the basic invention. An increase in the total number of claims should correlate, however, with the effort a firm puts into the patenting process, and this effort should increase in response to pressures for immediate and quantifiable results.

We do not use measures of originality and generality because their correspondence to exploration and exploitation remains unclear. The measures calculate the spread of classes covered by forward and backward citations, however, they do not take history into account; the spread may be novel and unique, or it may be old and common. For example, a patent may be measured as original because it cites other patents across a wide variety of classes, yet that citation pattern may have already appeared on any number of patents. Additional pragmatic issues make the measure unattractive: 1) it is only calculated for the NBER sample, 2) any calculation relies upon the concordance of classes which changes as each new class is defined, and 3) the USPTO recently stopped using the US class system, hence it will be impossible to update the measure going forward. Unreported regressions available from the first author show no significant effect of board transition on the average of a firm's patent scores of originality and generality, for the patents in the NBER sub sample. Individuallevel patent regressions similarly show no significant relationship.

4.2. Control variables

Following the extant literature, we control for a vector of firm characteristics that could confound the relation between board independence and a firm's innovative search and success. We compute all variables for firm *i* over its fiscal year t. Board size measures the number of board members as we want to insulate the effect of board independence from contemporary changes in the number of directors. Further, we found that the firms in our sample differ significantly in terms of R&D spending over total assets and firm size as measured by total assets-two variables that are naturally positively related to firms' innovation activities. In order to reduce the skewness in total assets we take the logarithm of total assets in all multivariate econometric analyses. In addition, we control for firm age (the number of years since the initial public offering date), as older firms may search in older technological areas. Moreover, leverage (long-term debt over total assets) and *capital expenditures* (scaled by total assets) account for financial constraints that are known to influence corporate innovation. Finally, Tobin's Q enters the regression to control for differences in growth opportunities.

4.3. Summary statistics

Table 1 presents summary statistics on the data set. The patenting activities of the firms in our sample show

⁵ Reflecting that a value of one indicates no change, the measure takes value one if no patent was applied for in a given year. All results presented below are robust to excluding non-patenting firms.

Summary statistics.

This table reports summary statistics of variables used in the study. Board size is the number of board members. Independent board is an indicator variable that indicates whether the majority of board members are independent. Top (1%) are the number of patents that fall into the 1% most cited patents within a given three-digit class and application year. Top 10% to 2% are the number of patents that fall into the 10% to 2% most cited patents within a given three-digit class and application year. Cited patents are the number of patents that received at least one citation but do not appear in the top 10% of the citation distribution. Uncited are the number of patents that were not cited. Self-citations are the number of cites to patents held by the same firm. Patents in new/known classes is the number of patents that are filed in classes where the given firm has filed no/at least one other patent beforehand. Tech. prox. is the technological proximity between the patents filed in year t to the existing patent portfolio held by the same firm up to year t-1, calculated according to laffe (1989). Further information on variable definitions and data sources provided in Sections 4.1 and 4.2.

Variable	Ν	Mean	Median	Sd	Min	Max
Patents	6107	53.78	3	243.36	0	5261
Citations	6107	573.70	5	3329.21	0	108,496
Claims	6107	1006.30	28.00	4718.74	0	88,533
Top 1%	6107	0.53	0	2.42	0	44
Top 10% to 2%	6107	5.06	0	25.18	0	660
Cited patents	6107	30.62	1	149.93	0	3512
Uncited patents	6107	18.13	1	98.77	0	4033
Back-citations	6107	1157.22	26	4851.25	0	101,943
Self-citations	6107	176.60	0	990.85	0	22,415
New classes	6107	1.28	0	3.88	0	227
Known classes	6107	52.50	2	242.47	0	5259
Tech. prox.	6107	0.54	0.68	0.41	0	1
Indep. board	6107	0.77	1	0.42	0	1
Board size	6107	9.23	9	2.52	3	21
log(Total assets)	6107	7.41	7.22	1.51	3.09	13.53
R&D /assets	6107	0.05	0.02	0.07	0	1.12
Age	6107	17.78	15	10.98	1	37
Cap. exp. /assets	6107	0.05	0.04	0.04	0	0.43
Leverage	6107	0.18	0.17	0.16	0	1.35
log(Q)	6107	1.23	1.04	0.85	-2.46	6.72

typical skewness with a mean of ~54 patents and a median of three patents. Related measures like the amount of R&D investment and citation-weighted patent counts reveal similar distributions and high concentrations among the most active firms. We calculated the number of patents that cite a given patent based on all US granted patents by April 2014. To control for secular trends in citation rates we employ time fixed effects that presumably affect all firms equally on average (see also Atanassov, 2013; and Hall, Jaffe, and Trajtenberg, 2001). 680 firms (85%) have applied for at least one patent during the sampling period. The average firm has filed 0.5 (5.0 patents) in the top 1% (10%) category, 18.1 patents that are never cited, and 30.6 that appear in the middle of the citation distribution. Similar to the number of cites received in the future, the number of backward citations is guite large on average with 1,157.2 cites (median 26). On average, 176.6 of those backward citations relate to patents that belong to the same firm (median 0). Further, 1.3 patents are filed in new-to-the-firm technology classes, while 52.5 are filed in known classes. The average technological proximity measure is 0.54.

Regarding other variables of interest, the average firm in our sample is 17.8 years old, has nine board members, a book value of assets of \$7 billion, an R&D to assets ratio of 5%, a leverage ratio of 18.2%, capital expenditures over total assets of 5.3%, and a Tobin's Q of 1.2. The Appendix displays a correlation table of all variables (Table 18).

4.4. Methodological remarks

In order to analyze how a transition to an independent board affects innovative search we follow the literature on corporate governance and innovation (e.g., Atanassov, 2013; He and Tian, 2013; Kortum and Lerner, 2000) and estimate the baseline model in Ordinary Least Square (OLS):

$$log(1 + patents_{i,t+1}) = \beta_0 + \beta_1 \cdot independentboard_{it} + \gamma \cdot Z_{it} + \theta_t + \alpha_i + \epsilon_{it},$$

where $patents_{i,t+1}$ is the number of eventually granted patents of firm *i* applied for in year *t*+1. In alternative regressions we will exchange the number of patents with our previously introduced measures of innovation that allow us to assess the firms' innovative search strategy in more detail.⁶ Our main explanatory variable of interest, independent board_{it}, is a dummy that indicates firms that have transitioned from a minority to a majority of independent board members in the year 2001 or later when regulatory changes became effective.⁷ Under the assumption that changes in patenting by firms that transitioned would have been comparable to changes in patenting by other firms in the absence of a transition, β_1 captures the effect of board independence on innovation by the affected firms.⁸ Z_{it} is a vector of the previously introduced firm characteristics, and year fixed effects θ_t control for changes in the macroeconomic environment and systematic changes in patenting activities over time. Our preferred specifications include firm fixed effects α_i that control for any unobserved firm heterogeneity that is time-invariant. Hence, we basically estimate a DiD model, where those firms that switch from a minority to a majority of independent directors on the board in 2001 or later are the 'treated firms', and all others are 'non-treated firms'. In order to unravel the influence of firm fixed effects in our regressions we also show alternative models with industry fixed effects, based on three-digit standard industry classification (SIC) industry dummies, instead of firm fixed effects. To stay

⁶ In case the dependent variable is a count, all results are robust to alternatively estimating Poisson models (not shown).

⁷ All results presented below are robust to alternatively taking the years 2000 or 2002 as the threshold value.

⁸ As can be seen in Fig. 1, not all firms transitioned from a friendly to an independent board at the same time, because directors were allowed to fulfill their contracts that were signed before the law change. In principle, this gives firms room for strategic choice that could confound our identification. Therefore, we checked whether the time between the law change and compliance is correlated with pre-SOX innovative activity of the firms in our sample. In order to test this, we first defined a variable that measures the years until the board actually changed from friendly to independent although SOX and other regulations were already active (2003). We found 17 firms with a one year lag, 14 with a two-year lag, and eight with a three-year lag. Then, we regressed time lag until compliance on firms' average amount of R&D, patents, and cites before 2001 (results are robust to taking 2000 or 2002 instead). The lack of significant correlation between compliance lags and pre-treatment innovative activity increases confidence that the estimation is not biased by systematic choice of more or less innovative firms to transition later or earlier.

within the DiD framework, we include a dummy variable that marks all treated firms in those regressions without firm fixed effects.

Identification hinges in all models upon the parallel trend assumption; treated and non-treated firms show similar trends in the dependent variable of interest in the absence of treatment. To increase our confidence in this assumption, we estimate the dynamics of the treatment effect, which provides evidence that the DiD estimator is not significantly different from zero in the absence of treatment.

Our estimation might still be biased, however, if other remaining cross-sectional heterogeneity of the firms in our sample change systematically with the transition to an independent board and our measures of innovative search. In order to minimize concerns in this regard, we further re-estimate all our models based on a balanced sample, where treated and non-treated firms are comparable in terms of key observable characteristics before 2002. To achieve a balanced sample we use Coarsened Exact Matching (CEM).⁹ CEM has several features that bound the degree of model dependence, reduce causal estimation error, bias, and inefficiency ((lacus, King, and Porro, 2012, 2011), for a similar application, see Azoulay, Zivin, and Wang (2010)). Based on CEM's coarsening function we match treated and non-treated firms on the joint distribution of firms' R&D spending over total assets, firm size as measured by the natural logarithm of total assets, the natural logarithm of Tobin's Q, boardsize, and 26 two-digit SIC industry code dummies. We took the average values of those variables over the years 2000 and 2001 as matching criteria to ensure highest comparability before treatment.¹⁰ Table 2 presents the differences in mean values of all control variables before and after the matching procedure.

Panels A and B of Table 2 show that treated firms in the full sample are on average a little smaller, invest less in R&D, and have a smaller board. Except with regard to R&D spending, the relative differences of the two firm groups appear small in magnitude. Both groups are not statistically significant with regard to the mean values of the other control variables that have not explicitly been included in the matching. In order to eliminate any statistically significant differences of observable firm characteristics, while keeping as many treated firms as possible in the sample, we ran CEM with the side condition to differentiate firms according to ten categories of R&D spending and three categories of firm size, board size, and Tobin's Q. Based on this procedure, four out of the 125 treated firms remain unmatched. For the remaining 121 treated firms, CEM selected 430 comparable firms, i.e., 158 incomparable firms are subsequently discarded from the analysis. Panel C of Table 2 shows that, after matching, there are no statistically significant differences between the treated and nontreated firms according to two sided *t*-tests. Although not necessary for a consistent DiD estimation, it is worthwhile

Table 2

CEM matching of treated and non-treated firms.

This table reports mean values of treated and non-treated observable firm characteristics, averaged over the years 2000 and 2001, before and after matching, based on the joint distribution of firms' R&D spending over total assets, firm size as measured by the natural logarithm of total assets, the natural logarithm of Tobin's Q, and board size. ***, **, * denote significance levels of 1%, 5%, and 10% of two-sided *t*-tests on the difference between mean values of Panels A and B, and Panels A and C, respectively.

Variable	Number of firms	Mean				
Panel A: Treated firms before matching						
log(Total assets) 125						
R&D /assets	125	0.04				
Age	125	2.45				
Leverage	125	0.18				
Cap. exp.	125	0.06				
log(Q)	125	1.34				
Board size	125	8.45				
Panel B: Non-treated firms	before matching					
log(Total assets)	588	7.33**				
R&D /assets	588	0.05*				
Age	588	2.43				
Leverage	588	0.20				
Cap. exp.	588	0.05				
log(Q)	588	1.25				
Board size	588	8.99**				
Panel C: Non-treated firms	after matching					
log(Total assets)	430	6.99				
R&D /assets	430	0.04				
Age	430	2.37				
Leverage	430	0.20				
Cap. exp.	430	0.05				
log(Q)	430	1.21				
Board size	430	8.56				

to mention that both groups do not differ in terms of the average amount of applied patents after matching.

While balancing the sample should improve identification (at least for firms that are similar to the treated firms), potential remaining differences in innovation trends might still have an influence on the estimation. Therefore, we also estimate models that allow for separate firm-specific linear trends in innovation before 2002, using the following specification:

 $log(1 + patents_{i,t+1}) = \beta_0 + \beta_1 \cdot independent board_{it}$

 $+\gamma \cdot Z_{it} + \delta \cdot firm_i \cdot pre2002_t \cdot t + \theta_t + \alpha_i + \epsilon_{it},$

where $pre2002_t$ equals one if the year of observation is 2001 or earlier.

Finally, in alternative specifications we further control for potential systematic changes in the influence of our controls on innovation after 2001, which may coincide with changes in board independence, by estimating:

 $log(1 + patents_{i,t+1}) = \beta_0 + \beta_1 \cdot independent \ board_{it} + \gamma \cdot Z_{it} + \delta \cdot firm_i \cdot pre2002_t \cdot t + \zeta \cdot Z_{it} \cdot post_t + \theta_t + \alpha_i + \epsilon_{it}.$

5. Results

We first present results on research and development spending, the number of patents, and the total number

⁹ In alternative models we balanced the sample based on propensity score matching, taking only the nearest neighbor of each treated firm as a control, and find qualitatively the same results.

¹⁰ The results are robust to taking all available observations before 2001 into account.

Independent boards and R&D.

The dependent variable is log(R&D). All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.822***	0.564***	0.601***	0.609***	0.602***
	(0.017)	(0.044)	(0.040)	(0.049)	(0.049)
log(Age)	-0.153***	0.002	-0.006	-0.017	-0.013
	(0.021)	(0.029)	(0.038)	(0.056)	(0.054)
Leverage	-0.562***	0.040	-0.085	-0.211	-0.462**
	(0.113)	(0.107)	(0.124)	(0.152)	(0.212)
Cap. exp.	0.753	0.562	0.542	0.378	0.820
	(0.616)	(0.351)	(0.391)	(0.431)	(0.518)
log(Q)	0.366***	-0.016	-0.015	-0.014	0.022
	(0.025)	(0.024)	(0.029)	(0.032)	(0.035)
Board size	0.024**	0.007	0.004	0.006	-0.004
	(0.009)	(0.008)	(0.011)	(0.013)	(0.014)
Independent board	0.071	-0.052	-0.057	-0.059	-0.043
	(0.090)	(0.055)	(0.056)	(0.064)	(0.061)
Observations	6107	6107	4414	4414	4414
R ²	0.733	0.256	0.254	0.450	0.508
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

of citations to and claims within a firm's patent portfolio. We then present measures of innovative search strategy, including a breakdown of the citation distribution, backward and self-citations, and movement into new classes and across technological distance.

5.1. R&D, patents, citation-weighted patents, and claims

Tables 3-6 estimate regressions of firms' R&D investments, the number of eventually granted patents applied for, the total number of citations made to the firm's patents, and the total number of claims contained within a firm's patent portfolio. Each table contains five specifications of the same model. Specification (a) is a standard OLS model with industry fixed effects, (b) is a standard firm fixed effects model, (c) is the same as (b) but estimated on the previously described balanced CEM sample, (d) adds trend controls, and (e) adds interaction terms of all controls with a post-SOX marker. For all models with firm fixed effects the R-squared values refer to the explained within-firm variance. The first model assesses potential changes in R&D investments after board independence changed, which might drive subsequent changes in patenting.¹¹ The next two models differentiate between a change in the number of patents and a change in cita-

Table 4

Independent boards and number of patents.

The dependent variable is the logarithm of one plus the number of eventually granted patents. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.767***	0.273***	0.284***	0.369***	0.425***
	(0.017)	(0.060)	(0.064)	(0.067)	(0.079)
R&D	5.561***	0.941*	0.842	0.711	0.835
	(0.568)	(0.517)	(0.668)	(0.713)	(0.896)
log(Age)	0.105***	0.068	0.000	0.004	-0.019
	(0.023)	(0.044)	(0.039)	(0.048)	(0.058)
Leverage	-0.468^{***}	-0.112	-0.094	-0.253	-0.250
	(0.123)	(0.176)	(0.196)	(0.188)	(0.212)
Cap. exp.	1.635***	0.147	0.127	0.321	0.325
	(0.490)	(0.484)	(0.518)	(0.522)	(0.561)
log(Q)	0.199***	0.057*	0.057	0.081**	0.066
	(0.027)	(0.034)	(0.037)	(0.040)	(0.041)
Board size	0.015	0.017	-0.003	-0.016	-0.012
	(0.010)	(0.014)	(0.016)	(0.015)	(0.017)
Independent board	0.308***	0.272***	0.215***	0.208**	0.198**
-	(0.083)	(0.079)	(0.080)	(0.087)	(0.087)
Observations	6107	6107	4414	4414	4414
R^2	0.571	0.207	0.176	0.410	0.414
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

tions to those patents. Cite-weighted patent counts (the total number of citations to a firm's patent portfolio) have been shown to correlate with financial value and patent renewals (Harhoff, Narin, Scherer, and Vopel, 1999; and Hall, Jaffe, and Trajtenberg, 2005). The last model estimates effects on the total number of claims in a firm's portfolio.

Table 3 illustrates that a transition to an independent board appears unrelated to the level of firms' R&D investments. In contrast, Tables 4 and 5 illustrate how patenting and total citations both increase. Reading across the models, the effect on patenting ranges between a 31% to 20% increase in the number of patents, and a 59% to 41% increase in total citations. Fig. 2 illustrates the dynamics of the latter two effects. For the graphs we defined dummy variables for the specific times before and after firms changed to an independent board. t_0 defines the year of the switch and serves as the baseline category, t_{n-1} defines the number of years before the switch, and t_{n+1} the years after the switch. Then, we ran regressions including these variables instead of the single dummy variable in the baseline model beforehand. As we still include year fixed effects, the coefficients represent the relative change in patenting per year that is attributable to the board change.

Table 6 illustrates that the number of claims in a firm's patents increases following a transition to an independent board. The effect on the number of claims ranges between

¹¹ Alternative regressions with R&D investments scaled by total assets reveal a significant positive effect only in specifications without firm fixed effects. Inclusion of controls for time-invariant firm heterogeneity leads to statistically insignificant results.

Independent boards and number of cite-weighted patents.

The dependent variable is the logarithm of one plus the number of citation-weighted patents. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.919***	0.329***	0.287***	0.326***	0.525***
R&D	(0.027) 7.750***	(0.090) 2.451***	(0.099) 2.671**	(0.115) 3.257***	(0.127) 4.779***
10 m(A ma)	(0.870)	(0.835)	(1.036)	(1.162)	(1.413)
log(Age)	(0.038)	(0.072)	(0.013)	(0.029)	-0.052 (0.090)
Leverage	-0.374*	0.110	0.323	0.191	0.162
Cap. exp.	(0.200) 2.636***	(0.261) 0.163	(0.301) 0.233	(0.304) 0.610	(0.388) 0.567
	(0.803)	(0.821)	(0.856)	(0.972)	(1.118)
log(Q)	0.351***	0.220***	0.237***	0.240***	0.272*** (0.086)
Board size	-0.002	-0.005	-0.031	-0.048*	-0.048
Independent board	(0.015) 0.472*** (0.146)	(0.021) 0.738*** (0.161)	(0.026) 0.599*** (0.169)	(0.027) 0.498** (0.214)	(0.032) 0.499** (0.212)
Observations	6107	6107	4414	4414	4414
R^2	0.505	0.316	0.284	0.445	0.454
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	NO NO	Yes No	Yes No	Yes Yes	Yes Yes
Controls * post-SOX	No	No	No	No	Yes

a 50% and 36% increase in the number of claims. Fig. 3 illustrates the dynamics.

The results are consistent with classic agency theory and our first hypothesis, suggesting that intensified monitoring leads to increased effort of the agent, which results in increased claims and patenting of inventions. That firms patent more, but do not spend significantly more on R&D, raises the question whether firms just work more efficiently or exploit extant knowledge at the expense of explorative innovation (models of patenting efficiency were not significant). Our second hypothesis proposes that increased board independence leads to a shift from explorative to more exploitative innovative activities. The following models illustrate a consistent shift towards exploitation but no clear signal of the effect on exploration.

5.2. The distribution of citations

Most recent research that uses patent data considers raw counts and total citations to the raw count patents, but less research considers the distribution of citations in careful detail. In this section we model the number of breakthrough, important, incremental, and failed inventions that a firm makes. These estimations are motivated by the argument that responding to increased oversight will increase

Table 6

Independent boards and number of claims.

The dependent variable is the logarithm of one plus the total number of claims of a patent portfolio. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	1.013***	0.380***	0.362***	0.512***	0.477***
	(0.030)	(0.099)	(0.114)	(0.132)	(0.146)
R&D	8.685***	1.326	1.362	1.486	0.580
	(0.981)	(0.950)	(1.182)	(1.300)	(1.368)
log(Age)	0.146***	0.030	-0.002	-0.001	-0.097
	(0.043)	(0.059)	(0.067)	(0.083)	(0.094)
Leverage	-0.291	0.188	0.230	-0.099	0.084
	(0.228)	(0.282)	(0.322)	(0.330)	(0.414)
Cap. exp.	1.448	-0.023	0.012	0.447	0.338
	(0.881)	(0.934)	(0.989)	(1.059)	(1.183)
log(Q)	0.280***	0.110*	0.152**	0.184**	0.172**
	(0.051)	(0.062)	(0.068)	(0.078)	(0.087)
Board size	0.000	0.002	-0.014	-0.037	-0.036
	(0.017)	(0.022)	(0.027)	(0.028)	(0.033)
Independent board	0.501***	0.488***	0.476***	0.365**	0.359**
	(0.153)	(0.137)	(0.142)	(0.178)	(0.178)
Observations	6107	6107	4414	4414	4414
R ²	0.466	0.133	0.119	0.304	0.307
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

tangible and countable but incremental patents at the expense of risky patents; such risky patents are more likely to fail completely or provide a breakthrough.

To model each of these four possible outcomes, we split the distribution into subcategories: (1) the number of patents that the firm invents that received cites within the highest percentile (top 1%) among all patents in the same three-digit patent class and application year, (2) the number of patents that received cites within the highest centile (10%) among all patents in the same three-digit patent class and application year but not including the top 1%. (3) the number of patents that received at least one citation (the median of the entire distribution is zero) but not including the top 10%, and (4) the number of patents that received no citation. Hence, the measures should be interpreted as 1) the number of breakthroughs, 2) the number of important patents, 3) the number of incremental patents that have small value, and 4) the number of patents that have little or no value to the firm. As an example, in the year 2000, IBM invented 4,367 patents, of which 24 were in the top 1% of their field, 360 in the top 10% but not including the top 1%, 3,374 with at least one cite but not in the top 1% or 10%, and 609 of which received no citations. Tables 7-10 present the corresponding results for each of the bins.



Fig. 2. Dynamics of independent board effect on patents and citations. These figures illustrate the effect of a change in board independence on patenting and citations over time. For the graphs we defined dummy variables for the time firms changed from a minority of independent board members to an independent board. t_0 indicates the year of the switch and serves as the reference category. t_{n-1} indicate the years before the switch, and t_{n+1} the corresponding years after the switch. Coefficients are taken from the last regression model of Section 4.4, but with the t_n dummies instead of the one dummy variable indicating a majority of independent board members.

Consistent with the models in Tables 4 and 5 we see a positive effect of board transitions on patenting and citation rates. The estimated effect is by far the most significant and largest-from 35% to 22%-for incremental patents that received at least one citation (but not in the top 10% of the distribution), while the estimated effect on particularly successful patents (top 1% or top 10%) is very small in magnitude and significant at p < 0.10 for only two out of ten regressions. Taking also into account that the effect on the number of unsuccessful patents (no cites) is most often statistically insignificant, the evidence is consistent with the argument that firms focus on less risky opportunities when the board becomes independent. Inclusion of a measure of backward citations weakens these effects further, implying that the increase in citations is mediated by movement of the firm into more crowded areas of technological search (models not shown but available from first author). In other words, the increase in citations may not correspond to an increase in patent value, rather, it may be an artifact of the exploitation strategy.



Fig. 3. Independent boards and number of claims. This figure illustrates the effect of a change in board independence on claims over time. For the graph we defined dummy variables for the time firms changed from a minority of independent board members to an independent board. t_0 indicates the year of the switch and serves as the reference category. t_{n-1} indicate the years before the switch, and t_{n+1} the corresponding years after the switch. Coefficients are taken from the last regression model of Section 4.4, but with the t_n dummies instead of the one dummy variable indicating a majority of independent board members.

Table 7

The number of breakthrough inventions: independent boards and top 1% patents.

The dependent variable is the logarithm of one plus the number of patents that fall in the top 1% percentile of the citation distribution within patent class and application year. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.166***	0.037**	0.056***	0.054***	0.037*
	(0.008)	(0.016)	(0.015)	(0.019)	(0.020)
R&D	0.724***	-0.092	-0.060	-0.045	-0.102
	(0.097)	(0.136)	(0.223)	(0.290)	(0.364)
log(Age)	0.036***	0.013	0.004	-0.002	-0.008
	(0.007)	(0.010)	(0.008)	(0.011)	(0.013)
Leverage	-0.198***	-0.049	-0.113***	-0.145^{**}	-0.110
	(0.035)	(0.042)	(0.043)	(0.058)	(0.070)
Cap. exp.	0.489***	-0.109	-0.094	-0.110	-0.049
	(0.150)	(0.118)	(0.107)	(0.130)	(0.152)
log(Q)	0.031***	-0.000	-0.015	-0.021	-0.023
	(0.009)	(0.011)	(0.012)	(0.015)	(0.015)
Board size	0.000	0.004	0.000	-0.002	0.002
	(0.003)	(0.003)	(0.004)	(0.004)	(0.005)
Independent board	0.027	0.043*	0.030	0.045*	0.041
	(0.027)	(0.024)	(0.025)	(0.027)	(0.026)
Observations	6107	6107	4414	4414	4414
R^2	0.312	0.009	0.014	0.179	0.182
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

The number of important inventions: independent boards and top 2–10% patents.

The dependent variable is the logarithm of one plus the number of patents that fall in the top 10% centile of the citation distribution within patent class and application year (excluding the top 1%). All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a)	(b)	(c)	(d)	(e)
	D/Se	D/Se	D/Se	D/Se	D/Se
log(Total assets)	0.389***	0.103***	0.113***	0.109***	0.058
	(0.014)	(0.030)	(0.032)	(0.035)	(0.038)
R&D	2.283***	0.218	-0.072	-0.120	-0.232
	(0.265)	(0.222)	(0.344)	(0.420)	(0.553)
log(Age)	0.072***	0.040**	0.027*	0.030	0.035
	(0.015)	(0.017)	(0.014)	(0.019)	(0.022)
Leverage	-0.300***	0.049	-0.046	-0.079	-0.064
	(0.076)	(0.072)	(0.083)	(0.095)	(0.110)
Cap. exp.	0.997***	-0.236	-0.196	-0.228	-0.068
	(0.330)	(0.207)	(0.219)	(0.228)	(0.277)
log(Q)	0.101***	0.030	0.026	0.034	0.028
	(0.018)	(0.019)	(0.022)	(0.030)	(0.028)
Board size	0.003	0.004	-0.004	-0.007	-0.001
	(0.007)	(0.006)	(0.007)	(0.008)	(0.009)
Independent board	0.069	0.064*	0.051	0.062	0.061
	(0.054)	(0.039)	(0.040)	(0.055)	(0.054)
Observations	6107	6107	4414	4414	4414
R^2	0.407	0.017	0.021	0.208	0.214
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

5.3. Self- and backward citations

In this section, we investigate more nuanced measures of search strategy. First, we focus on the number of citations that each patent makes to other patents. An increase in the number of backward citations reflects more prior art that must be specified in the patent application. This should correlate with innovative search in relatively better-known and mature technological areas. Second, we take the number of times a given patent cites other patents owned by the same company. More self-cites indicate constraining search within previously known areas of expertise while fewer self-citations indicate a broadening of innovative search or efforts to explore areas that are new to the firm. Tables 11 and 12 present the corresponding results and Fig. 4 illustrates the dynamics of the effects.¹² Firms that transition to independent boards increase backward and self-citations right after the transition and the effect remains for subsequent years.

The results presented in Tables 11 and 12 and Fig. 4 support the argument that firms with independent

Table 9

The number of incremental inventions: independent boards and cited patents, not in top 10%.

The dependent variable is the logarithm of one plus the number of patents that are cited but do not fall in the top 10% of the citation distribution. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
	Бузе	0/30	0/30	0/30	6/30
log(Total assets)	0.678***	0.268***	0.227***	0.251***	0.316***
	(0.018)	(0.055)	(0.057)	(0.062)	(0.069)
R&D	4.879***	1.123**	0.820	0.857	1.210
	(0.497)	(0.459)	(0.566)	(0.615)	(0.755)
log(Age)	0.097***	0.045	0.001	0.004	-0.024
	(0.023)	(0.034)	(0.032)	(0.041)	(0.048)
Leverage	-0.433***	-0.045	-0.031	-0.103	-0.064
	(0.116)	(0.148)	(0.157)	(0.162)	(0.189)
Cap. exp.	2.093***	0.284	0.407	0.553	0.544
	(0.481)	(0.401)	(0.419)	(0.455)	(0.519)
log(Q)	0.183***	0.091***	0.103***	0.097***	0.090**
0.1.5	(0.027)	(0.031)	(0.032)	(0.037)	(0.039)
Board size	0.004	0.009	-0.003	-0.016	-0.014
	(0.009)	(0.012)	(0.014)	(0.014)	(0.016)
Independent board	0.348***	0.339***	0.260***	0.229***	0.220***
•	(0.076)	(0.067)	(0.067)	(0.073)	(0.074)
Observations	6107	6107	4414	4414	4414
R^2	0.536	0.248	0.207	0.416	0.421
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

boards tend to narrow their innovative search towards known and mature technological areas. Regression coefficients imply an increase of 50% to 39% for backward citations and 39% to 26% for self-citations. The effect holds in patent-level regressions as well—it is not an artifact of increased patenting (see Appendix for robustness checks).

5.4. Technology classes

We now turn to the number of patents that are filed in USPTO classes previously unknown to the firm (the office classified all technologies into approximately 400 major classes). Unknown patent classes are defined as those in which a given firm has not been granted any patent back to 1976. The complement is the number of patents applied for in known classes. A more sophisticated measure of whether firms stay or deviate from known research areas is the technological proximity between the patents filed in year *t* and the existing patent portfolio held by the same firm up to year t-1 (Jaffe, 1989). Both measures use the original class in which the patent was granted (each year the USPTO changed the organization or concordance of the classifications).

Tables 13–15 present the corresponding regression results. Fig. 5 illustrates the dynamics of the effects on patents in known and unknown areas. As can be seen,

¹² Alternative regressions of non-self-citations reveal very similar results as estimated for the total number of backward citations.

The number of failed inventions: independent boards and patents without citations.

The dependent variable is the logarithm of one plus the number of patents that are not cited. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.635***	0.223***	0.278***	0.390***	0.299***
	(0.015)	(0.068)	(0.075)	(0.081)	(0.095)
R&D	3.953***	0.184	-0.206	-0.677	-1.452
	(0.433)	(0.557)	(0.868)	(0.990)	(1.322)
log(Age)	0.085***	0.071	-0.011	-0.004	-0.025
	(0.019)	(0.047)	(0.039)	(0.051)	(0.066)
Leverage	-0.418^{***}	-0.233	-0.299	-0.492^{**}	-0.273
	(0.103)	(0.175)	(0.209)	(0.217)	(0.244)
Cap. exp.	1.043***	-0.264	-0.304	-0.338	-0.343
	(0.400)	(0.472)	(0.513)	(0.532)	(0.568)
log(Q)	0.114***	0.003	-0.012	0.006	-0.025
	(0.023)	(0.035)	(0.040)	(0.044)	(0.043)
Board size	0.019**	0.028*	0.010	-0.000	0.003
	(0.008)	(0.015)	(0.015)	(0.015)	(0.018)
Independent board	0.167**	0.106	0.077	0.099	0.098
	(0.071)	(0.089)	(0.090)	(0.094)	(0.091)
Observations	6107	6107	4414	4414	4414
R^2	0.510	0.045	0.040	0.323	0.332
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

independent boards have an insignificant effect on exploration of new classes but a strong and significantly positive effect on search in previously patented classes—the number of patents in known classes increases by 32–20%. The Jaffe measure of technological proximity ranges from 25% to 29% but loses significance in the trend models.

6. Robustness checks and discussion

We first test robustness by adding controls for several alternative governance mechanisms that might confound the relationship between board independence and innovation. Second, we re-estimate all models using an instrumental variable (IV)-regression specification as introduced by Duchin, Matsusaka, and Oguzhan (2010). Third, we acknowledge limitations and close with a discussion of the mechanisms that might drive our results.

6.1. Governance provisions

Empirical research in corporate governance has considered a wide range of provisions that influence corporate behavior. If those factors simultaneously vary with board independence, missing controls could lead to an under- or overestimation of the independent board effect. We minimize this possibility by adding controls sequentially and finally estimating a full model. In order to economize on

Table 11

Independent boards and backward citations.

The dependent variable is the logarithm of one plus the number of backward citations. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	1.030***	0.399***	0.383***	0.556***	0.535***
DOD	(0.031)	(0.106)	(0.119)	(0.134)	(0.148)
R&D	8.023***	1.155	1.117	1.342	0.465
	(0.932)	(1.032)	(1.211)	(1.394)	(1.359)
log(Age)	0.133***	0.017	-0.022	0.009	-0.102
	(0.045)	(0.064)	(0.076)	(0.094)	(0.106)
Leverage	-0.231	0.213	0.251	-0.137	-0.052
	(0.238)	(0.304)	(0.356)	(0.363)	(0.439)
Cap. exp.	2.028**	0.085	0.044	0.382	0.515
	(0.915)	(0.958)	(1.007)	(1.099)	(1.225)
log(Q)	0.305***	0.127**	0.160**	0.188**	0.183**
	(0.052)	(0.063)	(0.070)	(0.083)	(0.091)
Board size	0.002	-0.002	-0.021	-0.041	-0.044
	(0.018)	(0.023)	(0.028)	(0.030)	(0.035)
Independent board	0.498***	0.479***	0.482***	0.389**	0.388**
-	(0.159)	(0.133)	(0.139)	(0.173)	(0.174)
Observations	6107	6107	4414	4414	4414
R ²	0.450	0.115	0.106	0.295	0.298
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

the size of the robustness check table we focus on five outcome variables: (1) Future citations received, (2) Breakthrough patents that fall in the Top 1% category, (3) Incremental patents that received at least one citation but do not fall into the Top 10% category, (4) Self-citations, (5) Patents filed in technology classes that are known to the firm. Reported results come from the previously introduced regression models with firm fixed effects and after matching.

We first consider CEO and board tenures, outside directorships, outside CEOs, and CEO shareholder voting rights. The sample shrinks because the IRRC collected data on those measures only from 1998 onwards. The tenure of the CEO and/or board members may influence investment decisions, through, for example, anticipated time horizons and payoffs to long-term investments. Hence, we add the years since the CEO has been appointed and the average time since appointment of all directors. Prior research has also shown that simultaneous outside board mandates can have an influence on the advice and control offered by the board (e.g., Field, Lowry, and Mkrtchyan, 2013) and that CEOs of other companies who simultaneously serve as monitoring directors influence corporate governance (Fahlenbrach, Low, and Stulz, 2010). The total number of simultaneous outside directorships of all board members and the number of outside CEOs on the board thus also enter the regressions (results are robust to

Independent boards and self-citations.

The dependent variable is the logarithm of one plus the number of selfcitations. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.833***	0.204***	0.160**	0.244***	0.244**
	(0.024)	(0.071)	(0.072)	(0.085)	(0.099)
R&D	5.728***	0.234	-0.104	-0.233	-0.585
	(0.636)	(0.671)	(0.835)	(1.056)	(1.106)
log(Age)	0.158***	0.075	0.013	0.030	0.000
	(0.031)	(0.047)	(0.045)	(0.060)	(0.071)
Leverage	-0.321*	0.004	-0.014	-0.275	-0.154
	(0.170)	(0.228)	(0.257)	(0.264)	(0.346)
Cap. exp.	3.488***	0.853	0.761	1.017	1.240
	(0.672)	(0.580)	(0.595)	(0.703)	(0.829)
log(Q)	0.269***	0.035	0.038	0.059	0.075
	(0.038)	(0.042)	(0.043)	(0.047)	(0.053)
Board size	0.018	0.023	0.010	-0.013	-0.017
	(0.013)	(0.015)	(0.018)	(0.018)	(0.022)
Independent board	0.389***	0.359***	0.284***	0.260***	0.262***
-	(0.109)	(0.080)	(0.081)	(0.096)	(0.096)
Observations	6107	6107	4414	4414	4414
R ²	0.469	0.088	0.061	0.285	0.286
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

taking the average number of outside directorships and CEOs, respectively; not reported). Table 16, Panel A, reports the results. Though the number of outside directorships of all board members is significantly and positively correlated with breakthrough innovations, none of the newly added board characteristics change the previously identified effect of the switch to an independent board (there is also no impact on the effect of switching on other dependent variables, both here and below, results omitted for brevity).

Next, we add a control variable for inside firm ownership which is proxied by the percentage of shareholder voting rights of the CEO (results are robust to alternatively taking the cumulative voting rights of all board members). Data come again from IRRC. Table 16, Panel B, illustrates how inside ownership itself is not significantly correlated with any innovation measure. The size of the marginal effect of board independence decreases by 10% to 15%, though the statistical significance of the effect remains at the 1% level.

Large shareholders have incentives to actively control executives (Shleifer and Vishny, 1997). Outside control may have changed with board independence as well, which might confound our previous estimations. Following Atanassov (2013), we control for large shareholder presence by adding a dummy variable that marks firms with at least one non-executive holding of at least 5% of eq-



Fig. 4. Dynamics of independent board effect on backward and selfcitations. *Notes:* These figures illustrate the effect of a change in board independence on backward and self-citations over time. For the graphs we defined dummy variables for the time firms changed from a minority of independent board members to an independent board. t_0 indicates the year of the switch and serves as the reference category. t_{n-1} indicate the years before the switch, and t_{n+1} the corresponding years after the switch. Coefficients are taken from the last regression in 4.4, but with the t_n dummies instead of the one dummy variable indicating a majority of independent board members.

uity (results are robust to taking higher threshold levels). Data come from Thomson Financial. Table 16, Panel C, reports the results. The presence of a blockholder appears to have no effect on the first four dependent variables, and a small and weakly significant negative effect on patents filed in known technology classes. Again, the independent board effect remains qualitatively unaltered.

The fifth robustness check adds a measure of the strength of the shareholders' rights using the G-Index of Gompers, Metrick, and Ishii (2003). The G-Index combines 24 corporate governance provisions that influence shareholder rights. It ranges from 1 to 24, where the lowest values indicate the strongest rights, and vice versa. Some of those combined provisions are particularly likely to have changed simultaneously with board independence. Hence, we add the following four indicators for firms that have: (1) a staggered board, where only a proportion of the directors can be replaced each year; (2) a poison pill that gives shareholders special rights to prevent hostile

Independent boards and patents in known classes.

The dependent variable is the logarithm of one plus the number of patents filed in classes where the given firm had already at least one other patent filed any previous year. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autcorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a)	(b)	(c)	(d)	(e)
	b/se	b/se	b/se	b/se	b/se
log(Total assets)	0.779***	0.275***	0.290***	0.368***	0.431***
R&D	5.718***	0.989*	0.743	0.534	0.649
log(Age)	(0.572)	(0.527)	(0.678)	(0.732)	(0.943)
	0.107***	0.061	-0.018	-0.020	-0.038
Leverage	(0.023)	(0.046)	(0.040)	(0.050)	(0.059)
	-0.521***	-0.206	-0.217	-0.383**	-0.390*
Can eyn	(0.124) 1.622***	(0.179)	(0.200)	(0.191) 0.122	(0.221)
	(0.500)	(0.471)	(0.510)	(0.501)	(0.553)
$\log(Q)$	0.214***	0.059*	0.064*	0.091**	0.080*
	(0.028)	(0.034)	(0.037)	(0.040)	(0.042)
Board size	0.018*	0.023	0.006	-0.005	-0.001
	(0.010)	(0.014)	(0.016)	(0.015)	(0.017)
Independent board	0.323***	0.289***	0.231***	0.209**	0.198**
	(0.082)	(0.079)	(0.079)	(0.083)	(0.083)
Observations	6107	6107	4414	4414	4414
R ²	0.572	0.184	0.153	0.401	0.406
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

takeover attempts; (3) cumulative voting where shareholders are allowed to combine their voting rights in a way that multiplies their rights with the number of directors to be elected (the only governance provision that is supposed to increase shareholder control), and (4) director indemnification where bylaws or the corporate charter can indemnify officers and directors from certain legal expenses and judgments resulting from lawsuits. Table 16, Panel D reports the results including the control for the G-Index, and Panel E reports the results including the specific governance provisions. None of the governance controls are significantly correlated with any of our dependent variables of interest. Again, the estimated effect of the independent board change remains qualitatively unaltered.

Finally, we re-estimate all our models jointly including all of the previously mentioned additional controls. Results are reported in Panel F of Table 16. Although this exercise reduces the amount of observations and controls for corporate governance provisions of all sorts, the effect of a switch to an independent board remains statistically and economically significant in models 1, 3, 4, and 5; consistent with our former results it stays insignificant in model 2. The sizes of the marginal effects decrease by up to 28% (in case of future citations). That only a few newly added controls show statistically significant influences themselves

Table 14

Independent boards and patents in unknown classes.

The dependent variable is the logarithm of one plus the number of patents filed in classes where the given firm had no other patent filed in any previous year. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a)	(b)	(c)	(d)	(e)
	b/se	b/se	b/se	b/se	b/se
log(Total assets)	0.171***	0.096***	0.103***	0.101**	0.177***
	(0.008)	(0.032)	(0.037)	(0.042)	(0.045)
R&D	0.994***	0.329	0.508	0.516	1.039**
	(0.175)	(0.270)	(0.379)	(0.472)	(0.494)
log(Age)	-0.002	0.018	0.025	0.032	0.021
	(0.010)	(0.019)	(0.021)	(0.025)	(0.029)
Leverage	-0.091^{*}	0.099	0.104	0.099	-0.011
	(0.055)	(0.086)	(0.098)	(0.115)	(0.134)
Cap. exp.	0.958***	0.916***	0.802**	0.866**	0.975**
	(0.239)	(0.303)	(0.322)	(0.389)	(0.443)
log(Q)	0.062***	0.064***	0.053***	0.026	0.012
	(0.012)	(0.019)	(0.020)	(0.025)	(0.027)
Board size	0.002	0.004	-0.004	-0.009	-0.012
	(0.005)	(0.008)	(0.011)	(0.012)	(0.012)
Independent board	0.088**	0.055	0.023	0.037	0.036
	(0.036)	(0.045)	(0.044)	(0.053)	(0.052)
Observations	6107	6107	4414	4414	4414
R ²	0.319	0.134	0.115	0.284	0.291
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

is partly related to the inclusion of firm fixed effects in all specifications, which limits investigations of these variables' effects, for example, the governance provisions, that exhibit low variation over time.

6.2. 2SLS estimation

Most similar to the present study, Duchin, Matsusaka, and Oguzhan (2010) estimated a 2-stage least square (2SLS) model to investigate if outside directors influence firms' market valuation. Instead of focusing on the switch to an independent board that happened during the sampling period, Duchin, Matsusaka, and Oguzhan (2010) focus on the percentage change of independent directors between 2000 and 2005 that was the result of regulatory changes that forced firms to appoint a majority of independent board members. Re-assembling their estimation strategy allows to test the robustness of our results with regard to the chosen empirical model as well as to assess the effect of a percentage change in independent directors. We estimate the same empirical model but use our previously introduced standard control variables on the righthand side of the equation [taking the exact same controls as Duchin, Matsusaka, and Oguzhan (2010) reveals qualitatively similar results].

Independent boards and technological proximity.

Notes: The dependent variable is the technological proximity between the patents filed in year t to the existing patent portfolio held by the same firm up to year t-1, and is calculated according to Jaffe (1989). All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.451***	0.187**	0.224**	0.363***	0.306**
	(0.020)	(0.085)	(0.100)	(0.117)	(0.124)
R&D	5.353***	0.859	0.976	0.820	0.033
	(0.710)	(0.835)	(1.018)	(1.104)	(1.132)
log(Age)	0.085***	-0.005	-0.024	-0.021	-0.049
	(0.030)	(0.052)	(0.063)	(0.077)	(0.081)
Leverage	-0.020	-0.138	-0.075	-0.448	-0.472
	(0.172)	(0.263)	(0.313)	(0.344)	(0.428)
Cap. exp.	-0.280	-0.179	-0.362	-0.228	-0.161
	(0.663)	(0.856)	(0.937)	(0.944)	(1.055)
log(Q)	0.177***	0.038	0.061	0.100	0.097
	(0.036)	(0.048)	(0.057)	(0.066)	(0.074)
Board size	0.014	0.023	0.011	0.010	0.022
	(0.012)	(0.018)	(0.022)	(0.025)	(0.030)
Independent board	0.247**	0.255**	0.289**	0.177	0.169
	(0.115)	(0.120)	(0.126)	(0.138)	(0.140)
Observations	6107	6107	4414	4414	4414
R^2	0.369	0.118	0.112	0.292	0.294
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

Specifically, we regress the change in our innovation variables from 2000 to 2005 on the change in the percentage of outside directors during the same time. By differencing the dependent and main independent variables of interest, all unobserved time-invariant effects are effectively taken into account. Causality is established by instrumenting the change in independent directors with an indicator that marks all firms that were non-compliant in the year 2000 with the regulatory changes that forced firms to assign a majority of independent board members. Both stages of the empirical model contain the previously introduced control variables in levels of the year 2000 to control for the initial firm characteristics that may have influenced changes in our innovation variables. Table 17 reports the results for (1) Future citations received, (2) Breakthrough patents that fall in the Top 1% category, (3) Incremental patents that received at least one citation but do not fall into the Top 10% category, (4) Self-citations, (5) Patents filed in technology classes that are known to the firm.

Results are mainly consistent with prior estimations; changes in the percentage of independent directors lead to significant increases of future citations (also number of patents, claims, and backward citations, not reported), in-



Fig. 5. Dynamics of independent board effect on patents in known and unknown classes These figures illustrate the effect of a change in board independence on patents filed in known and unknown classes over time. For the graphs we defined dummy variables for the time firms changed from a minority of independent board members to an independent board. t_0 indicates the year of the switch and serves as the reference category. t_{n-1} indicate the years before the switch, and t_{n+1} the corresponding years after the switch. Coefficients are taken from the last regression of Section 4.4, but with the t_n dummies instead of the one dummy variable indicating a majority of independent board members.

cremental innovations, and patents in known technological areas. Self-citations increase insignificantly, however, and there is a weakly significant, small positive effect on breakthrough inventions. The sizes of the marginal effects are also consistent with our previous models. Considering an increase in the percentage of independent directors of 8.5%, the 2SLS models predict a corresponding increase in future citations of 33.3%, an increase of incremental innovations of 16.4%, and an increase of patents with known technologies of 11.0% (total patents 14.1%, *p*-value < 0.01; claims 31.5%, *p*-value < 0.01; backward cites 36.1%, p-value < 0.01). Overall, the robustness check largely supports our main findings from the DiD models.

6.3. Discussion

The results consistently describe a shift towards innovative exploitation for firms that transition to an

Robustness checks.

This table reports the results of separate robustness checks in each panel. The dependent variables are the logarithm of one plus the total number of citations received (model A), the logarithm of one plus the number of patents in the top 1% of the citation distribution per year and tech class (model B), the logarithm of one plus the number of patents that received at least one citation but do not fall in the top 10% of the citation distribution per year and tech class (model C), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of patents filed in known technology classes (model E). All specifications are estimated after matching and include firm and time fixed effects and control variables as introduced in Section 4.2, not shown. Independent variables are introduced in Section 6.1. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) Citations	(b)	(c) Cited no ten 10%	(d) Salf citos	(e) Known toch
	b/se	b/se	b/se	b/se	b/se
Panel A: Board characteristics	,	,	,	,	,
Director tenure	0.003	0.002	0.001	-0.001	0.002
	(0.018)	(0.002)	(0.010)	(0.012)	(0.011)
CEO tenure	0.022	-0.002	0.005	-0.007	0.025
	(0.041)	(0.006)	(0.022)	(0.030)	(0.026)
Outside directorships	0.088	0.026**	0.047	0.059	0.061
	(0.069)	(0.012)	(0.036)	(0.049)	(0.046)
Outside CEOs	0.067	0.013	0.030	-0.042	0.023
	(0.078)	(0.013)	(0.043)	(0.055)	(0.048)
Independent board	0.517***	0.033	0.254***	0.289***	0.237***
	(0.121)	(0.026)	(0.065)	(0.074)	(0.075)
Panel B: Inside control					
Inside control	0.048	0.004	0.019	-0.002	0.029
	(0.037)	(0.008)	(0.018)	(0.022)	(0.020)
Independent board	0.469***	0.048	0.200***	0.233***	0.214***
	(0.107)	(0.032)	(0.056)	(0.067)	(0.068)
Panel C: Outside control					
Blockholder	-0.072	-0.002	-0.052	-0.003	-0.087*
	(0.081)	(0.011)	(0.040)	(0.049)	(0.049)
Independent board	0.539***	0.030	0.262***	0.284***	0.234***
	(0.119)	(0.025)	(0.067)	(0.081)	(0.079)
Panel D: Corporate governance in	dex				
G-Index	0.053	-0.057	0.191	0.043	0.284
	(0.341)	(0.042)	(0.176)	(0.209)	(0.218)
Independent board	0.548***	0.031	0.272***	0.275***	0.242***
	(0.115)	(0.026)	(0.065)	(0.078)	(0.078)
Panel E: Corporate governance pr	ovisions				
Staggered board	0.377	0.004	0.192	0.260	0.157
	(0.272)	(0.026)	(0.158)	(0.248)	(0.177)
Poison pill	-0.040	-0.026	0.056	-0.022	0.016
	(0.139)	(0.022)	(0.084)	(0.117)	(0.097)
Cumulative voting	-0.413	-0.037	-0.095	0.004	-0.053
	(0.396)	(0.035)	(0.161)	(0.155)	(0.170)
Director indemnification	0.118	0.018	0.094	0.119	-0.005
	(0.230)	(0.084)	(0.145)	(0.149)	(0.225)
Independent board	0.537***	0.030	0.266***	0.268***	0.238***
	(0.114)	(0.026)	(0.064)	(0.077)	(0.078)
Panel F: All controls					
Independent board	0.382***	0.042	0.193***	0.222***	0.235***
-	(0.119)	(0.035)	(0.060)	(0.068)	(0.070)

independent board. They do not provide consistent evidence for any influence on exploration; there are positive but weakly significant increases in the tail of the citation distribution and no impact on patenting in new classes. Here we summarize robustness checks reported in the Appendix, explore potential mechanisms that could accomplish the shift towards exploitation, and discuss why independent boards (and managers) might have less influence on exploration.

Firms which transition to independent boards patent more; this raises the concern that the increased backward and self-citation results might simply be artifacts of the increased patenting. To rule out this possibility, we estimate regressions of backward and self-citations per patent.

Robustness checks.

The dependent variables are the logarithm of one plus the total number of citations received (model A), the logarithm of one plus the number of patents in the top 1% of the citation distribution per year and tech class (model B), the logarithm of one plus the number of patents that received at least one citation but do not fall in the top 10% of the citation distribution per year and tech class (model C), the logarithm of one plus the total number of self-citations (model D), the logarithm of one plus the total number of patents filed in known technology classes (model E). Alndependent directors are the predicted changes in the percentage of independent directors between 2000 and 2005. The first stage includs all control variables introduced in Section 4.2, industry dummies, and, as an instrument, a dichotomous variable that marks all firms that did not comply with SOX in the year 2000. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) ∆Citations b/se	(b) ∆Top 1% b/se	(c) Δ Cited, no top 10% b/se	(d) ∆Self-cites b/se	(e) ∆Known tech b/se
log(Total assets)	-0.181***	0.011	-0.143***	-0.069	-0.123**
	(0.061)	(0.013)	(0.027)	(0.049)	(0.045)
R&D	-5.273*	-0.232	-1.561	-0.021	-1.598
	(2.680)	(0.349)	(1.869)	(2.456)	(2.064)
log(Age)	0.106	0.005	0.088*	0.185***	0.115*
	(0.111)	(0.021)	(0.044)	(0.058)	(0.061)
Leverage	0.343	0.160	0.267	-0.404	-0.280
	(0.696)	(0.128)	(0.413)	(0.432)	(0.462)
Board size	-0.100***	-0.020**	-0.054**	-0.075	-0.053
	(0.030)	(0.009)	(0.023)	(0.052)	(0.032)
Δ Independent directors	0.039***	0.006*	0.019**	0.007	0.013***
(predicted values)	(0.009)	(0.003)	(0.007)	(0.006)	(0.004)
Observations	416	416	416	416	416
R ²	0.215	0.074	0.201	0.098	0.139

As can be seen in Tables 19 and 20 in the Appendix, the proportion of backward and self-citations also increases for firms which transition to independent boards. Effects sizes range from 22-18% for backward citations and 16-14% for self-citations. We also investigated the coefficient of variance of citations to firms that undergo the transition to independent boards. While the results were not significant on a yearly basis, an aggregation of the four years following the transition demonstrated a significant decrease; consistent with a shift to exploitation and thinning of the tails, citations to firms with independent boards become less variable after the transition. We found no evidence that the transition to independent boards influences innovative efficiency (Cohen, Diether, and Malloy, 2013); there is no statistically significant result for the regression of patents per R&D investment and citations per R&D are positive but lose significance in the trend models.

Table 21 shows that our results are more pronounced for firms with high research and development spending and stock. Board independence appears to have a stronger impact on firms for which innovation is more important (high R&D), probably because for those firms the tension between exploration/exploitation is more significant and boards need to be more concerned about their innovation strategies. If innovation is less important (low R&D), there is probably less board involvement in innovation and thus our results are less pronounced.

As proposed in the introduction, multiple mechanisms could cause a firm whose board becomes independent to shift towards exploitation. For example, managers may shirk less and work harder in response to greater oversight, take less risk out of career concerns, respond to advice, or search less because they fear an independent board will constrain future flexibility. Ruling out one versus another mechanism empirically remains difficult, as they imply similar predictions and probably co-exist in practice. Nonetheless, the split sample tests described below remain consistent with models where greater oversight results in increased effort and risk aversion.

We split the sample into firms with high and low managerial entrenchment, using the index of Bebchuk, Cohen, and Ferrell (2009). This entrenchment or "eindex" indicates how many corporate governance provisions are in place that shield a manager from getting fired, e.g., poison pills, golden parachutes, and supermajority requirements for mergers and charter amendments.¹³ Table 22 shows that the effects of independent boards are consistently stronger for firms with high managerial entrenchment. Managers in firms with low entrenchment index are already subject to career concerns and takeover pressures, even before the board becomes independent. Therefore, the transition to an independent board does not have much impact on these managers. It is for entrenched managers that transitioning into independent boards can trigger career concerns and greater exploitation.

Most likely, many mechanisms play overlapping and possibly complementary roles in explaining the shift towards exploitation. The aggregate evidence is most

¹³ The E-Index is given by Bebchuk, Cohen, and Ferrell (2009) for all equal years and is fairly stable over time. In order to keep the sample size as large as possible we imputed with the lagged value where the E-Index was missing; if the lagged value was missing we took the forward value.

consistent, however, with increased oversight and career concerns mechanisms; the shift towards exploitation when a firm's board becomes independent is most likely due to a combination of greater managerial effort and an increased aversion to innovative risk.

We do not see an increase in exploitation at the expense of breakthroughs, indeed, we see almost no impact of board transition on breakthroughs at all. This might imply that exploration and exploitation are not ends of a continuum; rather, they could be orthogonal and possibly complementary strategies, especially for firms with large research portfolios and relatively independent research teams. Alternately, exploitation could also lead to breakthroughs, through greater focus and reliance on deep expertise (Jones, 2009). Particularly with the increasing "burden of knowledge," an exploration strategy necessarily implies a more shallow understanding of an area simply because it is new and unfamiliar. This could be an even more severe handicap for fast moving fields. A lack of significance might also simply result from lack of statistical power; breakthroughs are by definition rare events.

It is likely, however, that exploitation crowds out exploration in general (March, 1991). To begin with, the first-order goal of boards and managers is value creation for shareholders, not scientific or technical discovery. It is also likely that boards and managers have less fundamental ability to influence exploration: it is simply easier to manage and organize extant and tangible possibilities than to encourage breakthrough creativity. By its nature, exploration also takes longer to measure and appropriate. To explore this hypothesis with the current data set, we ran lagged outcomes of later years, but found no significant differences. Such delayed and uncertain payback is less attractive to managers and boards. It may also be more difficult to appropriate explorative innovation, as it requires a stable workforce and patient investment; this hypothesis could be explored with patent citation diffusion models.

These arguments raise the possibility that the contradictory results in the literature on governance and innovation are caused by the conflation of two successful innovation strategies. For example, this research and that of Lerner, Sorensen, and Stromberg (2011) illustrate an exploitation strategy; in contrast, Chemmanur and Tian (2016) illustrate an exploration strategy (as evidenced by the increase in citation variance). Both probably lead to an increase in firm value, though the mechanisms, risk, and payoff vary greatly. As a further example, are the nonmonotonic results of Sapra, Subramanian, and Subramanian (2014) due to differing search strategies? Does higher citation come from focus and exploitation on the strong governance end and search and exploration on the weak governance end? Similarly, does the weak governance result come from the tails of the distribution and greater volatility in patent citations, consistent with an exploration strategy? Some of the conflicting results in the governance and innovation literature might be profitably revisited with more nuanced measures and an effort to better identify the particular mechanisms and strategies that result in increased patenting and citations.

7. Conclusion

We proposed that firms which undergo a transition to more independent boards increase exploitation of previously successful areas of expertise. We argued that the shift towards exploitation results from stronger board oversight which increases both managerial effort and risk aversion. Supporting evidence came from the regulatory changes of Sarbanes-Oxley; firms that transition to more independent boards invent more but less explorative patents. On average these patents also receive more citations, though the citations occur to patents in the middle of the distribution and not to breakthrough or completely failed patents. Furthermore, the increase in cites is due partly to an increase in claims within each patent as well as movement into more crowded areas of technology. This implies that the increase in citations is partly due to a more thorough patenting of extant portfolios and an artifact of the citation norms in more crowded fields. Firms that transition also patent more heavily in technology classes of their current portfolio; they do not patent more in new classes. The effects are more pronounced for research-intensive firms and those that score high on measures of managerial entrenchment. Speaking to the larger literature on governance and innovation, our results indicate that strengthened governance improves innovation performance along existing trajectories, without harming the probability of a breakthrough.

We offered more nuanced and easily calculated patent measures that enable greater insight into the search and innovation process. These measures highlight the importance of differentiating between the greater productivity of exploitation and the riskier search of exploration. The results indicate that firms can increase their patent counts and even future citations to those patents—through exploitation of their existing portfolios. Further work should differentiate, both theoretically and empirically, between greater and focused effort and riskier search; it should not assume that an increase in patent counts or citations implies an increase in risk-taking and creativity.

Independent boards appear to move firms towards innovative exploitation and have little impact on exploration, but what is best for performance? Other research has found mixed evidence for the impact of independent boards on overall performance (see, e.g., Duchin, Matsusaka, and Oguzhan, 2010; Nguyen and Nielsen, 2010; Adams, Hermalin, and Weisbach, 2010). Lack of exploration may cause long-term obsolescence and competency traps, but where is the optimal tradeoff? Can large and diverse firms avoid the stark tradeoff, by developing portfolios that simultaneously explore and exploit? These are topics for future research.

Appendix. Robustness checks

Tables 19 and 20 report the average number of backward and self-citations, demonstrating that firms' exploitation is not an artifact of greater patenting. Tables 21 and 22 report the split samples by R&D and entrenchment, illustrating stronger effects for research-intensive firms and those with more entrenched managers.

Cross-correlations.

This table reports pair-wise correlations of all variables used in the study. Board size is the number of board members. Independent board is an indicator variable that indicates whether the majority of board members are independent. Top (1%) are patents that fall into the 1% most cited patents within a given three-digit class and application year. Top 10–2% are patents that fall into the 10–2% most cited patents within a given three-digit class and application year. Cited patents that received at least one citation but do not appear in the top 10%. Uncited are the number of patents that were not cited. Self-citations are the number of cites to patents held by the same firm. Patents in new/known classes is the number of patents that are filed in classes where the given firm has filed no/at least one other patent beforehand. Tech. prox. is the technological proximity between the patents filed in year *t* to the existing patent portfolio held by the same firm up to year *t*–1, and is calculated according to Jaffe (1989). Further information on variable definitions and data sources are provided in Section 4.2.

Variables	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20
1 Patents	1.00																			
2 Citations	0.74	1.00																		
3 Claims	0.98	0.76	1.00																	
4 Top1%	0.84	0.63	0.82	1.00																
5 Top10% to 2%	0.92	0.63	0.89	0.91	1.00															
6 Cited patents	0.92	0.89	0.94	0.74	0.78	1.00														
7 Uncited patents	0.83	0.32	0.76	0.70	0.83	0.55	1.00													
8 Back-citations	0.89	0.71	0.92	0.77	0.81	0.86	0.68	1.00												
9 Self-citations	0.84	0.74	0.85	0.69	0.73	0.83	0.61	0.89	1.00											
10 New classes	0.24	0.23	0.24	0.20	0.19	0.26	0.15	0.25	0.17	1.00										
11 Known classes	1.00	0.74	0.98	0.84	0.92	0.92	0.83	0.89	0.84	0.22	1.00									
12 Tech. prox.	0.19	0.15	0.18	0.19	0.17	0.18	0.15	0.20	0.16	0.15	0.19	1.00								
13 Indep. board	0.06	0.03	0.05	0.06	0.06	0.05	0.06	0.06	0.04	0.02	0.06	0.00	1.00							
14 Board size	0.14	0.09	0.12	0.14	0.13	0.12	0.13	0.14	0.13	0.07	0.14	0.01	0.10	1.00						
15 log(Total assets)	0.35	0.24	0.33	0.33	0.33	0.31	0.30	0.34	0.28	0.17	0.35	0.10	0.15	0.56	1.00					
16 R&D /assets	0.07	0.08	0.07	0.06	0.06	0.08	0.05	0.06	0.04	0.08	0.07	0.17	0.02	-0.25	-0.22	1.00				
17 Age	0.11	0.05	0.10	0.11	0.11	0.08	0.11	0.08	0.09	-0.02	0.11	-0.03	0.07	0.28	0.23	-0.18	1.00			
18 Cap. exp. /assets	0.04	0.08	0.04	0.01	0.01	0.07	0.00	0.04	0.05	0.07	0.04	0.04	-0.05	-0.01	-0.02	0.01	-0.05	1.00		
19 Leverage	-0.06	-0.06	-0.07	-0.06	-0.05	-0.06	-0.05	-0.06	-0.03	-0.04	-0.06	-0.06	0.03	0.16	0.14	-0.18	-0.03	0.01	1.00	
20 log(Q)	0.10	0.12	0.10	0.10	0.09	0.10	0.06	0.11	0.07	0.08	0.10	0.20	-0.00	-0.20	-0.07	0.44	-0.13	-0.07	-0.31	1.00

Table 19

Independent boards and average number of backward cites.

The dependent variable is the logarithm of one plus the average number of backward citations per patent. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.041**	0.054	0.021	0.002	-0.008
R&D	(0.017) 0.012	(0.050) -0.025	(0.065) -0.281	(0.082) -0.230	(0.088) -0.443
log(Age)	(0.378) -0.038	$(0.486) \\ -0.004$	(0.710) -0.007	(0.877) 0.053	(0.873) 0.002
Leverage	(0.025)	(0.035)	(0.040)	(0.052)	(0.062)
Levelage	(0.130)	(0.143)	(0.175)	(0.224)	(0.259)
Cap. exp.	1.068* (0.545)	0.623 (0.464)	0.527 (0.512)	-0.144 (0.737)	-0.175 (0.774)
$\log(Q)$	0.093***	0.028	0.036	0.033	0.043
Board size	-0.017	-0.005	-0.004	-0.014	-0.020
Independent board	(0.010) 0.217** (0.104)	(0.011) 0.154* (0.082)	(0.016) 0.197** (0.084)	(0.019) 0.179* (0.102)	(0.025) 0.181* (0.105)
	(0.104)	(0.002)	(0.004)	(0.102)	(0.105)
Observations R ²	3888 0.214	3888 0.022	2630 0.018	2630 0.269	2630 0.273
Year fixed effects Firm fixed effects	Yes No	Yes Yes	Yes Yes	Yes Yes	Yes Yes
Trend control Controls * post-SOX	No No	No No	No No	Yes No	Yes Yes
1					

Table 20

Independent boards and average number of self-cites.

The dependent variable is the logarithm of one plus the average number of self-citations per patent. All explanatory variables are lagged by one period. Specification (a) includes untabulated three-digit SIC industry dummies and a dummy that marks all treated firms. Independent board is a dummy that indicates firms after they switched from a minority of independent board members to a majority of independent board members in 2001 or later. Control variables are defined in Section 4.2. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a) b/se	(b) b/se	(c) b/se	(d) b/se	(e) b/se
log(Total assets)	0.086***	-0.046	-0.041	0.009	0.005
	(0.011)	(0.036)	(0.047)	(0.063)	(0.067)
R&D	0.654**	-0.347	-0.514	-0.641	-0.841
	(0.254)	(0.313)	(0.493)	(0.597)	(0.526)
log(Age)	0.035**	0.024	0.009	0.007	0.007
	(0.015)	(0.022)	(0.027)	(0.034)	(0.035)
Leverage	0.160*	-0.039	-0.144	-0.241	-0.165
	(0.086)	(0.119)	(0.149)	(0.194)	(0.245)
Cap. exp.	1.356***	0.488	0.464	0.419	0.504
	(0.331)	(0.350)	(0.403)	(0.582)	(0.616)
log(Q)	0.105***	-0.009	-0.028	-0.001	0.017
	(0.016)	(0.018)	(0.022)	(0.030)	(0.035)
Board size	0.007	0.014**	0.011	0.001	-0.008
	(0.006)	(0.007)	(0.009)	(0.011)	(0.013)
Independent board	0.156**	0.166***	0.167***	0.143**	0.144**
	(0.065)	(0.049)	(0.051)	(0.067)	(0.066)
Observations	3888	3888	2630	2630	2630
R ²	0.228	0.023	0.027	0.237	0.240
Year fixed effects	Yes	Yes	Yes	Yes	Yes
Firm fixed effects	No	Yes	Yes	Yes	Yes
Trend control	No	No	No	Yes	Yes
Controls * post-SOX	No	No	No	No	Yes

Median split, high and low R&D stock, matched sample.

The sample is split at the median value of the firms' R&D stock, calculated with a depreciation rate of 15%, using the perpetual inventory method introduced by Hall, Jaffe, and Trajtenberg (2005). Coefficients stem from two separate estimations based on each subsample. The dependent variables are the logarithm of one plus the total number of citations received (model a), the logarithm of one plus the number of patents in the top 1% of the citation distribution per year and tech class (model b), the logarithm of one plus the number of patents that received at least one citation but do not fall in the top 10% of the citation distribution per year and tech class (model c), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of patents filed in known technology classes (model e). All specifications include time fixed effects and control variables as introduced in Section 4.2, not shown. Independent board is a dummy that indicates firms after they switched from a minority of independent board members in 2001 or later. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a)	(b)	(c)	(d)	(e)
	Citations	Top 1%	Cited, no top 10%	Self-cites	Known tech
	b/se	b/se	b/se	b/se	b/se
Independent board	0.644***	0.065	0.322***	0.351**	0.280**
(High R&D stock)	(0.198)	(0.054)	(0.123)	(0.146)	(0.142)
Independent board	0.252**	-0.013	0.113**	0.105	0.083
(Low R&D stock)	(0.120)	(0.008)	(0.055)	(0.068)	(0.076)
Observations	2207	2207	2207	2207	2207
Firm fixed effects	Yes	Yes	Yes	Yes	Yes

Table 22

Median split, high and low entrenchment, matched sample.

The sample is split at the median value of the top managers' degree of entrenchment measured by the E-Index of Bebchuk, Cohen, and Ferrell (2009). Missing values of the E-Index cause a slightly smaller sample. Coefficients stem from two separate estimations based on each subsample. The dependent variables are the logarithm of one plus the total number of citations received (model a), the logarithm of one plus the number of patents in the top 1% of the citation distribution per year and tech class (model b), the logarithm of one plus the total number of self-citations (model d), the logarithm of one plus the total number of patents filed in known technology classes (model e). All specifications include time fixed effects and control variables as introduced in Section 4.2, not shown. Independent board is a dummy that indicates firms after they switched from a minority of independent board members in 2001 or later. Heteroskedasticity-robust standard errors that account for autocorrelation at the firm level are reported in parentheses. Coefficients: *** Significant at 1% level, ** Significant at 5% level, * Significant at 10% level.

	(a)	(b)	(c)	(d)	(e)
	Citations	Top 1%	Cited, no top 10%	Self-cites	Known tech
	b/se	b/se	b/se	b/se	b/se
Independent board	0.731***	0.003	0.430***	0.426***	0.475***
(High E-Index)	(0.203)	(0.021)	(0.108)	(0.107)	(0.124)
Independent board	0.418***	0.064*	0.148	0.161	0.020
(Low E-Index)	(0.154)	(0.038)	(0.090)	(0.117)	(0.105)
Observations	2145	2145	2145	2145	2145
Firm fixed effects	Yes	Yes	Yes	Yes	Yes

References

- Adams, R.B., Ferreira, D., 2007. A theory of friendly boards. Journal of Finance 62, 217–250.
- Adams, R.B., Hermalin, B.E., Weisbach, M.S., 2010. The role of boards of directors in corporate governance: a conceptual framework and survey. Journal of Economic Literature 48, 58–107.
- Aghion, P., van Reenen, J.M., Zingales, L., 2013. Innovation and institutional ownership. American Economic Review 103, 277–304.
- Ahern, A.R., Dittmar, A.K., 2012. The changing of the boards: the impact on firm valuation of mandated female board representation. Quarterly Journal of Economics 127, 137–197.
- Atanassov, J., 2013. Do hostile takeovers stifle innovation? evidence from antitakeover legislation and corporate patenting. Journal of Finance 68, 1097–1131.
- Atanassov, J., 2016. Arm's length financing and innovation: evidence from publicly traded firms. Management Science 62, 128–155.
- Azoulay, P., Graff Zivin, J.S., Manso, G., 2011. Incentives and creativity: evidence from the academic life sciences. RAND Journal of Economics 42 (3), 527–554.
- Azoulay, P., Zivin, J., Wang, J., 2010. Superstar extinction. Quarterly Journal of Economics 125, 549–589.
- Balsmeier, B., Chesebro, T., Fierro, G., Johnson, K., Kaulagi, A., Li, G., Lueck, S., O'Reagan, D., Yeh, B., Zang, G., Fleming, L., 2016. Machine learning and natural language processing on the patent corpus: data, tools, and new measures. Working paper, Fung Institute for Engineering Leadership: http://funginstitute.berkeley.edu/wp-content/ uploads/2016/11/Machine_learning_and_natural_language_processing_ on_the_patent_corpus.pdf http://www.funginstitute.berkeley.edu/ sites/default/files/weeklydbig0.09.pdf

- Bebchuk, L., Cohen, A., Ferrell, A., 2009. What matters in corporate governance? Review of Financial Studies 22, 783–827.
- Bernstein, S., 2014. Does going public affect innovation? The Journal of Finance 70, 1365–1403.
- Borokhovich, K.A., Parrino, R.P., Trapani, T., 1996. Outside directors and CEO selection. Journal of Financial and Quantitative Analysis 31, 337–355.
- Brickley, J.A., Coles, J.L., Terry, R.L., 1994. Outside directors and the adoption of poison pills. Journal of Financial Economics 35, 371–390.
- Brochet, F., Srinivasan, S., 2014. Accountability of independent directors: evidence from firms subject to securities litigation. Journal of Financial Economics 111, 430–449.
- Byrd, J.W., Hickman, K.A., 1992. Do outside directors monitor managers? evidence from tender offer bids. Journal of Financial Economics 32, 195–221.
- Byun, S., Oh, J., Xia, H., 2015. Deterring "creative" innovation: a potential negative externality of technology spillovers. Unpublished working Paper, University of Texas, Dallas
- Cerqueiro, G., Hegde, D., Penas, M., Seamans, R., 2015. Debtor rights, credit supply, and innovation. working paper
- Chemmanur, T., Tian, X., 2016. Do anti-takeover provisions spur corporate innovation?. Unpublished working paper, Boston College and Indiana University
- Chen, G., Gao, H., Hsu, P., Li, K., 2015. Ownership transition, managerial short-termism, and exploratory versus exploitative innovation strategy. Unpublished working paper, University of British Columbia
- Cohen, L., Diether, K., Malloy, C., 2013. Misvaluing innovation. Review of Financial Studies 26, 635–666. doi:10.1093/rfs/hhs183. Online first

- Coles, J.L., Daniel, N.D., Naveen, L., 2008. Boards: does one size fit all? Journal of Financial Economics 87, 329–356.
- Core, J.E., Holthausen, R.W., Larcker, D.F., 1999. Corporate governance, chief executive officer compensation, and firm performance. Journal of Financial Economics 51, 371–406.
- Cotter, J.F., Shivdasani, A., Zenner, M., 1997. Do independent directors enhance target shareholder wealth during tender offers? Journal of Financial Economics 43, 195–218.
- Denis, D.J., Sarin, A., 1999. Ownership and board structures in publicly traded corporations. Journal of Financial Economics 52, 187–223.
- Duchin, R., Matsusaka, J.G., Oguzhan, O., 2010. When are outside directors effective? Journal of Financial Economics 96, 195–214.
- Fahlenbrach, R., Low, A., Stulz, R.M., 2010. Why do firms appoint CEOs as outside directors? Journal of Financial Economics 97, 12–32.
- Faleye, O., Hoitash, R., Hoitash, U., 2011. The costs of intense board monitoring. Journal of Financial Economics 101, 160–181.
- Fama, E.F., Jensen, C.M., 1983. Separation of ownership and control. The Journal of Law & Economics 26, 301–325.
- Field, L., Lowry, M., Mkrtchyan, A., 2013. Are busy boards detrimental? Journal of Financial Economics 109, 63–82.
- Gompers, P.A., Metrick, A., Ishii, J.L., 2003. Corporate governance and equity prices. The Quarterly Journal of Economics 118, 107–155.
- Gonzalez-Uribe, J., Xu, M., 2016. CEO contract horizon and innovation. Unpublished working paper, London School of Economics
- Hall, B.H., Jaffe, A., Trajtenberg, M., 2001. The NBER patent citations data file: lessons, insights, and methodological tools. NBER Working Paper No. 8498
- Hall, B.H., Jaffe, A., Trajtenberg, M., 2005. Market value and patent citations. RAND Journal of Economics 16–38.
- Harhoff, D., Narin, F., Scherer, F.M., Vopel, K., 1999. Citation frequency and the value of patented inventions. Review of Economics and Statistics 81, 511–515.
- Harris, M., Raviv, A., 1978. Some results on incentive contracts with applications to education, insurance, and law enforcement. American Economic Review 68, 20–30.
- He, J., Tian, X., 2013. The dark side of analyst coverage: the case of innovation. Journal of Financial Economics 109 (3), 856–878.
- Holmstrom, B., 1979. Moral hazard and observability. Bell Journal of Economics 10, 74–91.
- Holmstrom, B., Milgrom, P., 1991. Multitask principal agent analyses: incentive contracts, asset ownership, and job design. Journal of Law, Economics, and Organization 7, 24–52.
- Hsu, P.-H., Tian, X., Xu, Y., 2014. Financial development and innovation: cross-country evidence. Journal of Financial Economics 112, 116–135.
- Iacus, S.M., King, G., Porro, G., 2011. Multivariate matching methods that are monotonic imbalance bounding. Journal of the American Statistical Association 106, 345–361.
- Iacus, S.M., King, G., Porro, G., 2012. Causal inference without balance checking: coarsened exact matching. Journal of Political Analysis 20, 1–24.

- Jaffe, A.B., 1989. Characterizing the "technological position" of firms, with application to quantifying technological opportunity and research spillovers. Research Policy 18, 87–97.
- Jensen, M.C., 1986. Agency costs of free cash flow, corporate finance, and takeovers. American Economic Review 76, 323–329.
- Jones, B.F., 2009. The burden of knowledge and the 'death of the renaissance man': is innovation getting harder? Review of Economic Studies 76, 283–317.
- Kang, J., Liu, W., Low, A., Zhang, L., 2014. Friendly boards and innovation. Unpublished working paper, Nanyang Technological University
- Knyazeva, A., Knyazeva, D., Masulis, R.W., 2013. The supply of corporate directors and board independence. Review of Financial Studies 26 (6), 1561–1605.
- Kortum, S., Lerner, J., 2000. Assessing the contribution of venture capital to innovation. RAND Journal of Economics 31, 674–692.
- Lanjouw, J., Schankerman, M., 2004. Patent quality and research productivity: measuring innovation with multiple indicators. The Economic Journal Volume 114, 441–465.
- Lerner, J., Seru, A., 2014. The use and abuse of patent data. Unpublished working paper, Harvard Business School
- Lerner, J., Sorensen, M., Stromberg, P., 2011. Private equity and long-run investment: the case of innovation. Journal of Finance 66, 445–477.
- Linck, J., Netter, J., Yang, T., 2008. The determinants of board structure. Journal of Financial Economics 87, 308–328.
- Manso, G., 2011. Motivating innovation. Journal of Finance 66 (5), 1823–1860.
- March, J., 1991. Exploration and exploitation in organizational learning. Organization Science 2, 71–87.
- Masulis, R.W., Mobbs, S., 2014. Independent director incentives: where do talented directors spend their limited time and energy? Journal of Financial Economics 111, 406–429.
- Nguyen, B.D., Nielsen, K.M., 2010. The value of independent directors: evidence from sudden deaths. Journal of Financial Economics 98, 550–567.
- Sapra, H., Subramanian, A., Subramanian, K.V., 2014. Corporate governance and innovation: theory and evidence. Journal of Financial and Quantitative Analysis 49, 957–1003.
- Shleifer, A., Vishny, R., 1997. A survey of corporate governance. Journal of Finance 52, 737–783.
- Sorensen, J., Stuart, T., 2000. Aging, obsolescence and organizational innovation. Administrative Science Quarterly 45, 81–112.
- Stein, J., 1989. Efficient capital markets, inefficient firms: a model of myopic corporate behavior. The Quarterly Journal of Economics 104, 655–669.
- Stiglitz, J., Weiss, A., 1983. Incentive effects of termination: applications to the credit and labor markets. American Economic Review 73, 912–927.
- Weisbach, M.S., 1988. Outside directors and CEO turnover. Journal of Financial Economics 20, 431–460.
- Williamson, O.E., 1983. Credible commitments: using hostages to support exchange. The American Economic Review 73, 519–540.